The campus grading regulations prohibit regarding of final exams. Therefore, read all parts of questions carefully and answer fully. Maximum number of points is 70.
I. Consider the following two-equation model in which the $y$ s are endogenous and the $x$ s are exogenous:

$$
\begin{aligned}
& \text { (1) } y_{1}=\alpha_{1} y_{2}+\alpha_{2} x_{1}+u \\
& \text { (2) } y_{2}=\beta_{1} y_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+v
\end{aligned}
$$

Ia ( 5 points) Check the order condition for each equation and state whether the condition is satisfied for exact or over identification. Explain your answers.

In the first equation, two variables are missing. This is one more than is needed and hence the equation is overidentified. The second equation has one variable missing and is exactly identified.
$\mathbf{I b}(10$ points) Explicitly derive the reduced forms for the two endogenous variables. You need not try to solve backwards for the original parameters.

Substituting for $y_{1}$ from the first equation into the second we get,

$$
y_{2}=\beta_{1}\left(\alpha_{1} y_{2}+\alpha_{2} x_{1}+u\right)+\beta_{2} x_{2}+\beta_{3} x_{3}+v
$$

This gives $\quad y_{2}\left(1-\alpha_{1} \beta_{1}\right)=\beta_{1} \alpha_{2} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+$ error. Dividing by $1-\alpha_{1} \beta_{1}$, we get

$$
\begin{aligned}
y_{2} & =\frac{\beta_{1} \alpha_{2}}{1-\alpha_{1} \beta_{1}} x_{1}+\frac{\beta_{2}}{1-\alpha_{1} \beta_{1}} x_{2}+\frac{\beta_{3}}{1-\alpha_{1} \beta_{1}} x_{3}+\text { error } \\
& =\pi_{1} x_{1}+\pi_{2} x_{2}+\pi_{3} x_{3}+\text { error }
\end{aligned}
$$

Substituting this into the first equation, we have

$$
\begin{aligned}
y_{1} & =\alpha_{1}\left(\pi_{1} x_{1}+\pi_{2} x_{2}+\pi_{3} x_{3}\right)+\alpha_{2} x_{1}+\text { error } \\
& =\left(\alpha_{1} \pi_{1}+\alpha\right) x_{1}+\alpha_{1} \pi_{2} x_{2}+\alpha_{1} \pi_{3} x_{3}+\text { error } \\
& =\mu_{1} x_{1}+\mu_{2} x_{2}+\mu_{3} x_{3}+\text { error }
\end{aligned}
$$

Ic ( $5+5$ points) Describe carefully how you would apply the two -stage least squares to each of the equations.

## Equation (1)

In the first stage, regress $y_{2}$ against $x_{1}, x_{2}$, and $x_{3}$, and save $\hat{y}_{2}$. Next regress $y_{1}$ against $\hat{y}_{2}$ and $x_{1}$, with no constant term.

## Equation (2)

In the first stage, regress $y_{1}$ against $x_{1}, x_{2}$, and $x_{3}$, and save $\hat{y}_{1}$. Next regress $y_{2}$ against $\hat{y}_{1}, x_{2}$, and $x_{3}$, with no constant term.
II.

The Justice Department hires you to relate the probability of conviction ( $p_{t}$ ) of an individual accused of a violent crime as a function of the number of previous convictions (CONVICT) for any crime, race (RACE), income (INCOME), and age (AGE). You have data on these variables and on $Y_{t}$, which takes the value 1 with probability $p_{t}$ if convicted, and 0 with probability $1-p_{t}$ if not convicted. You decide to use the Linear Probability Model

$$
\mathbf{Y}_{\mathbf{t}}=\beta_{1}+\beta_{2} \mathbf{C O N V I C T}_{\mathbf{t}}+\beta_{3} \mathbf{R A C E}_{\mathbf{t}}+\beta_{4} \mathbf{I N C O M E}_{t}+\beta_{5} \mathbf{A G E}_{\mathbf{t}}+\mathbf{u}_{\mathbf{t}}
$$

IIa. (15 points) Treating this as a binomial process rather than as a normal process, carefully show first that for $E\left(\mathbf{u}_{\mathbf{t}}\right)$ to be zero, $\mathbf{p}_{t}=\beta_{1}+\beta_{2}$ CONVICT $_{t}+\beta_{3}$ RACE $_{t}+$ $\beta_{4}$ INCOME $_{t}+\beta_{5}$ AGE $_{\mathbf{t}}$. Next show that the variance of $\mathbf{u}_{\mathbf{t}}$ is given by $\sigma_{t}^{2}=\mathbf{p}_{\mathbf{t}}\left(1-\mathbf{p}_{\mathbf{t}}\right)$.

When $Y_{t}=1$, we have $u_{t}=1-\beta_{1}-\beta_{2}$ CONVICT $_{t}-\beta_{3}$ RACE $_{t}-\beta_{4}$ INCOME $_{t}-\beta_{5} A G E_{t}$, with probability $p_{t}$. When $Y_{t}=0, u_{t}=-\beta_{1}-\beta_{2}$ CONVICT $_{t}-\beta_{3}$ RACE $_{t}-\beta_{4}$ INCOME $_{t}-$ $\beta_{5} A G E_{t}$, with probability $1-p_{t}$.

We have $0=E\left(u_{t}\right)=p_{t}\left(1-\beta_{1}-\beta_{2}\right.$ CONVICT $_{t}-\beta_{3}$ RACE $_{t}-\beta_{4}$ INCOME $\left._{t}-\beta_{5} A G E_{t}\right)-$ $\left(1-p_{t}\right)\left(\beta_{1}+\beta_{2}\right.$ CONVICT $_{t}+\beta_{3}$ RACE $_{t}+\beta_{4}$ INCOME $\left._{t}+\beta_{5} A G E_{t}\right)$

$$
=p_{t}-\beta_{1}-\beta_{2} \text { CONVICT }_{t}-\beta_{3} \text { RACE }_{t}-\beta_{4} \text { INCOME }_{t}-\beta_{5} A G E_{t}=0
$$

Solving for $p_{t}$ we get $p_{t}=\beta_{1}+\beta_{2}$ CONVICT $_{t}+\beta_{3}$ RACE $_{t}+\beta_{4}$ INCOME $_{t}+\beta_{5} A G E_{t}$.

We see that $u_{t}$ takes the value $1-p_{t}$ with probability $p_{t}$ and the value $-p_{t}$ with probability $1-p_{t}$. Therefore, $\sigma_{t}^{2}=E\left(u_{t}^{2}\right)=\left(1-p_{t}\right)^{2} p_{t}+p_{t}^{2}\left(1-p_{t}\right)=p_{t}\left(1-p_{t}\right)\left(1-p_{t}+p_{t}\right)=p_{t}\left(1-p_{t}\right)$.

IIb (10 points) Describe in full detail how the weighted least squares method can be used to estimate the $\beta$ s. Ignore any problems that might arise in this procedure.

First regress $Y_{t}$ against a constant, $\operatorname{CONVICT}_{t}, R A C E_{t}, \mathrm{INCOME}_{t}$, and $A G E_{t}$, and save $\hat{Y}_{t}$ which is also $\hat{p}_{t}$. Next compute the variance as $\hat{\sigma}_{t}^{2}=\hat{Y}_{t}\left(1-\hat{Y}_{t}\right)$. Then compute the weight $w_{t}=\frac{1}{\sqrt{\hat{\sigma}_{t}^{2}}}$. Finally regress $w_{t} Y_{t}$ against $w_{t}, w_{t}$ CONVICT $_{t}, w_{t} R A C E_{t}, w_{t} I N C O M E_{t}$, and $w_{t} A G E_{t}$.

## IIc (5 points) Explain what the drawbacks of this method are that make the approach unpopular.

First of all, when you regress $Y_{t}$ against a constant, CONVICT $_{t}, R A C E_{t}, I N C O M E ~_{t}$, and $A G E_{t}$, and save $\hat{Y}_{t}$, there is no guarantee that $0<\hat{Y}_{t}<1$. Second after getting the WLS estimators, there is no guarantee that $\hat{Y}_{t}$ obtained using the new estimates will give $\hat{Y}_{t}$ between 0 and 1 .

## III. (15 points) Describe in full how the Probit Model can be used for the model in II, instead of the linear probability function.

For the probit model, we assume that there is a response function $Y_{t}^{*}=\beta_{1}+\beta_{2}$ CONVICT $_{t}+$ $\beta_{3}$ RACE $_{t}+\beta_{4}$ INCOME $_{t}+\beta_{5} A G E_{t}+u_{t}$. We assume that $u_{t}$ is not binomial but normal with mean 0 and variance $\sigma_{t}^{2}$. Thus, $u_{t} / \sigma_{t}$ is the standard $N(0,1) . \quad Y_{t}=1$, if $Y_{t}^{*}>0$, and $Y_{t}=0$, if $Y_{t}^{*} \leq 0$. If $Z=u_{t} / \sigma_{t}$ is the standard normal and $F(z)$ is $P(Z \leq z)$, the cumulative distribution function of the standard normal, then we have

$$
\begin{aligned}
& P\left(Y_{t}=0\right)=P\left(Y_{t}^{*} \leq 0\right)=P\left(u_{t} \leq=-\beta_{1}-\beta_{2} \operatorname{CONVICT}_{t}-\beta_{3} \text { RACE }_{t}-\beta_{4} \text { INCOME }_{t}-\right. \\
& \left.\beta_{5} A G E_{t}\right)=P\left[u_{t} / \sigma \leq(1 / \sigma)\left(-\beta_{1}-\beta_{2} \text { CONVICT }_{t}-\beta_{3} R A C E_{t}-\beta_{4} I N C O M E ~_{t}-\right.\right. \\
& \left.\left.\beta_{5} A G E_{t}\right)\right]
\end{aligned}
$$

$$
=F\left[\left(1 / \sigma_{t}\right)\left(-\beta_{1}-\beta_{2} \text { CONVICT }_{t}-\beta_{3} R A C E_{t}-\beta_{4} I N C O M E ~_{t}-\beta_{5} A G E_{t}\right)\right]
$$

Call the express inside the function $F[], Q$. Then $P\left(Y_{t}=0\right)=F(Q)$.
Using the same procedure, we have $P\left(Y_{t}=1\right)=P\left(Y_{t}^{*}>0\right)=1-P\left(Y_{t}^{*} \leq 0\right)=1-F(Q)$. The above is a for a typical person. To get the full likelihood function, we multiply these for the two groups of accused, those for whom $Y_{t}=1$ and those for whom $Y_{t}=0$. The full likelihood function is therefore given by

$$
L=\prod_{Y_{t}=0} F(Q) \prod_{Y_{t}=1}[1-F(Q)]
$$

where $\Pi$ denotes the product of terms. Lis a function of the unknown $\beta$ s and $\sigma$. The parameters are estimated by maximizing this highly nonlinear function with respect to the $\beta s$ and $\sigma$. Many standard regression programs, including GRETL, carry out this procedure.

