Econ 120A Spring 2003 Ramu Ramanathan Answers to Exam #2

## I.

Let X and Y be two random variables, with means  $\mu_x$  and  $\mu_y$ ,  $Var(X) = \sigma_x^2 = E(X^2) - \mu_x^2$ ,  $Var(Y) = \sigma_y^2 = E(Y^2) - \mu_y^2$ , and  $Cov(X, Y) = \sigma_{XY}$ . Now make the transformations U = X + Y, and V = X - Y.

(a) (3 points) Derive E(U) and E(V) in terms of  $\mu_X$  and  $\mu_Y$ .

$$E(U) = E(X+Y) = E(X) + E(Y) = \mu_X + \mu_Y$$
  
$$E(U) = E(X-Y) = E(Y) - E(Y) = ii = ii$$

$$E(V) = E(X - Y) = E(X) - E(Y) = \mu_X - \mu_Y$$

(b) (5 points) Derive E(UV) in terms of  $\mu_X$ ,  $\mu_Y$ ,  $\sigma_X^2$  and  $\sigma_Y^2$ . [*Hint*: you want the expected value of U times V. Use the definitions at the top of the page.]

$$E(UV) = E[(X+Y)(X-Y)] = E(X^{2} - Y^{2}) = E(X^{2}) - E(Y^{2})$$
  
Since  $\sigma_{X}^{2} = E(X^{2}) - \mu_{X}^{2}$ , we have,  $E(X^{2}) = \sigma_{X}^{2} + \mu_{X}^{2}$ , similarly for  $E(Y^{2})$ .  
 $E(UV) = \sigma_{X}^{2} + \mu_{X}^{2} - (\sigma_{Y}^{2} + \mu_{Y}^{2})$ 

(c) (6 points) Derive Cov (*U*, *V*) in terms of  $\sigma_X^2$  and  $\sigma_Y^2$  only.

$$Cov (U, V) = E(UV) - [E(U) E(V)]$$
  
=  $\sigma_X^2 + \mu_X^2 - (\sigma_Y^2 + \mu_Y^2) - (\mu_X + \mu_Y)(\mu_X - \mu_Y)$   
=  $\sigma_X^2 + \mu_X^2 - (\sigma_Y^2 + \mu_Y^2) - (\mu_X^2 - \mu_Y^2) = \sigma_X^2 - \sigma_Y^2$ 

(d) (4 points) What is the condition under which U and V will be uncorrelated (that is, have zero correlation) and why?

Since uncorrelated means Cov(X, Y) = 0, required condition is  $\sigma_X^2 = \sigma_Y^2$ .

II.

Let  $x_1, x_2, ..., x_n$  be a random sample drawn from a population with mean  $\mu$  and variance  $\sigma^2$ . In other words,  $E(x_i) = \mu$ , and  $Var(x_i) = \sigma^2$  for i = 1, 2, ..., n, and the x's are all independent of each other. Let  $\overline{x} = \frac{1}{n} \sum_{i=1}^n x_i$  be the sample mean.

(a) (4 points) Show that  $E(\bar{x}) = \mu$ .

$$E(\bar{x}) = E\left(\frac{1}{n}\sum_{i=1}^{n}x_{i}\right) = \frac{1}{n}E\left(\sum_{i=1}^{n}x_{i}\right) = \frac{1}{n}\sum_{i=1}^{n}E(x_{i}) = \frac{1}{n}n \ \mu = \mu$$

It is possible to show that Var  $(\bar{x}) = \frac{\sigma^2}{n}$ , but you need not prove it.

(b) (4 points) Let 
$$Z = \frac{\sqrt{n(\overline{x} - \mu)}}{\sigma}$$
. Show that  $E(Z) = 0$ .

$$E(Z) = E(\frac{\sqrt{n}(\bar{x}-\mu)}{\sigma}) = \frac{\sqrt{n}}{\sigma}E(\bar{x}-\mu) = 0 \text{ because } E(\bar{x}) = \mu.$$

(c) (4 points) Show that Var(Z) = 1.

$$Var(Z) = E[\left(\frac{\sqrt{n}(\bar{x}-\mu)}{\sigma}\right)^2] = \left(\frac{\sqrt{n}}{\sigma}\right)^2 E[(\bar{x}-\mu)^2]$$
$$= \frac{n}{\sigma^2} Var(\bar{x}) = \frac{n}{\sigma^2} \frac{\sigma^2}{n} = 1$$

III.

X and Y are random variables with means  $\mu_x$  and  $\mu_y$ , variances  $\sigma_x^2$  and  $\sigma_y^2$ , and covariance  $\sigma_{xy}$ , all fixed constants. Consider the new random variable  $U = (Y - \mu_y) - b(X - \mu_x)$ , where b is a fixed constant that I choose. It follows that  $U^2 = (Y - \mu_y)^2 - 2b(X - \mu_x)(Y - \mu_y) + b^2(X - \mu_x)^2$ .

[Hint: Do not expand this expression. Keep the terms grouped as they are.]

(a) (4 points) Derive the expected value of U.

$$E(U) = E[(Y - \mu_Y) - b(X - \mu_x)] = E(Y) - \mu_Y - bE(X - \mu_x)$$
  
=  $E(Y) - \mu_Y - b[E(X) - \mu_x)] = \mu_Y - \mu_Y - b[\mu_X - \mu_x)] = 0$ 

(b) (7 points) Derive  $\sigma_U^2$ , the variance of *U*, in terms of *b*,  $\sigma_X^2$ ,  $\sigma_Y^2$  and the covariance  $\sigma_{XY}$ .

$$\sigma_U^2 = E[U^2] - [E(U)]^2 = E[U^2] \text{ because the second term is zero.}$$
  
$$= E[(Y - \mu_Y)^2 - 2b(X - \mu_X)(Y - \mu_Y) + b^2(X - \mu_X)^2]$$
  
$$= E[(Y - \mu_Y)^2] - 2bE[(X - \mu_X)(Y - \mu_Y)] + b^2E[(X - \mu_X)^2]$$
  
$$= \sigma_Y^2 - 2b\sigma_{XY} + b^2\sigma_X^2$$

(c) (6 points) Suppose I want to choose *b* in order to minimize  $\sigma_U^2$  with respect to *b*, *treating everything else as fixed*. Derive the value of *b* that minimizes this variance. [*Hint*: Your answer will depend only on  $\sigma_X^2$  and  $\sigma_{XY}$ ].

$$\frac{\partial \sigma_U^2}{\partial b} = 0 = -2\sigma_{XY} + 2b\sigma_X^2 \quad or \quad b\sigma_X^2 = \sigma_{XY}$$

Solving for b, we get  $b = \sigma_{XY} / \sigma_X^2$ .

(d) (3 points) Be sure to check the second-order condition also.

$$\frac{\partial^2 \sigma_U^2}{\partial b^2} = 2 \sigma_X^2 \quad \text{which is positive and hence a minimum is attained.}$$