

Answer key for the final exam in 2020

1a.) $\hat{\beta} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}(\mathbf{X}'\mathbf{V}^{-1}\mathbf{y})$

$\hat{\sigma}^2 = (T - k)^{-1}(\mathbf{y} - \mathbf{X}\hat{\beta})' \mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\hat{\beta})$

$\tau = \hat{\beta}_1 / \sqrt{\hat{\sigma}^2 \mathbf{e}_1' (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1} \mathbf{e}_1}$ for \mathbf{e}_1 first column of \mathbf{I}_k

τ has an exact small-sample Student t distribution with $T - k$ degrees of freedom under the stated assumptions

1b.) $h = [\mathbf{g}(\hat{\beta})]' \left[\hat{\sigma}^2 \mathbf{H}(\beta_0) (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1} \mathbf{H}(\beta_0)' \right]^{-1} \mathbf{g}(\hat{\beta})$

$\mathbf{H}(\beta_0) = \frac{\partial \mathbf{g}}{\partial \beta'} \big|_{\beta=\beta_0}$

Treat h as approximately $\chi^2(2)$; this is only asymptotic approximation

1c.) (1) White does not do anything with off-diagonal elements of \mathbf{V} but above procedure does; (2) White adjustment to diagonal terms is inefficiently estimated and only approximate; (3) White uses OLS estimate \mathbf{b} which is inefficient.

2.) $\text{plim } b = [E(x_t^2)]^{-1} [E(x_t y_t)]$ where

$E(x_t y_t) = E[x_t(g(x_t) + u_t)] = E[x_t g(x_t)]$

3a.) need $r = 2$ since we want to estimate both β_0 and β_1 . The first element of the (2×1) vector \mathbf{x}_t would usually be a constant term

3b.) Assuming that a constant term is first element of \mathbf{x}_t , regress z_t on \mathbf{x}_t . Do OLS F -test that coefficients other than constant term are all zero (we don't want to test that all coefficients are zero, since this is basically just testing whether z_t and x_t have mean zero). If F -statistic is less than 10, instruments might be weak.

4a.) $\frac{\partial \ell}{\partial p} = \frac{\sum x_t}{p} - \frac{\sum (n-x_t)}{1-p} = \frac{[(1-p) \sum x_t] - [p \sum (n-x_t)]}{p(1-p)} = \frac{[\sum x_t] - pTn}{p(1-p)}$.

Setting this to zero gives $\hat{p}_{MLE} = (Tn)^{-1} \sum x_t$

4b.) Notice \hat{p}_{MLE} is n^{-1} times the sample mean of T i.i.d. variables, each with variance $np(1-p)$. Thus

$\sqrt{T}(\hat{p} - p_0) \xrightarrow{L} n^{-1}N(0, np(1-p)) \sim N(0, p(1-p)/n)$.

No additional assumptions needed.

4c.) $h(p, x_t) = \frac{\partial \ell_t}{\partial p} = \frac{x_t}{p} - \frac{(n-x_t)}{(1-p)}$

$E[h(p, x_t)] = \frac{np_0}{p} - \frac{(n-np_0)}{(1-p)} = \frac{(1-p)np_0 - pn(1-p_0)}{p(1-p)} = \frac{n(p_0-p)}{p(1-p)}$ which equals 0 at $p = p_0$ and not

zero at all other p

4di) Since $r = a = 1$, \hat{p}_{GMM} satisfies

$T^{-1} \sum h(\hat{p}, x_t) = 0$ or $\frac{T^{-1} \sum x_t}{\hat{p}} - \frac{T^{-1} \sum (n-x_t)}{1-\hat{p}} = 0$, same FOC as MLE. Thus $\hat{p}_{MLE} = \hat{p}_{GMM}$.

4dii) $\hat{\mathbf{D}}' = T^{-1} \sum \frac{\partial h(p, x_t)}{\partial p} \bigg|_{p=\hat{p}} = T^{-1} \left[-\frac{\sum x_t}{\hat{p}^2} - \frac{\sum (n-x_t)}{(1-\hat{p})^2} \right] = -\frac{n\hat{p}}{\hat{p}^2} - \frac{n(1-\hat{p})}{(1-\hat{p})^2} = -\frac{n}{\hat{p}} - \frac{n}{(1-\hat{p})}$

4diii) $\hat{\mathbf{D}}' \xrightarrow{p} -\frac{n}{p_0} - \frac{n}{(1-p_0)} = -\frac{n}{p_0(1-p_0)}$

4div) $\hat{\mathbf{S}} = T^{-1} \sum [h(\hat{p}, x_t)]^2 = T^{-1} \sum \left[\frac{x_t}{\hat{p}} - \frac{(n-x_t)}{(1-\hat{p})} \right]^2 = \frac{T^{-1}}{\hat{p}^2(1-\hat{p})^2} \sum (x_t - n\hat{p})^2$