

Answer key for the final in 2009

- 1.) a.) v
 b.) v-vi
 c.) v-vi
 d.) vii
 e.) none
 f.) none
- 2.) a.) $\hat{\beta} = \left(\sum_{t=1}^T \underline{x}_t \underline{x}_t' \right)^{-1} \left(\sum_{t=1}^T \underline{x}_t y_t \right)$ for $\underline{x}_t = (1, x_t)'$.
 b.) $\hat{\gamma} = S_{xy}/S_{xx}$
 c.)

$$\begin{aligned} \sum_{t=1}^T e_t^2 &= \sum_{t=1}^T [(y_t - \bar{y}) - \hat{\gamma}(x_t - \bar{x})]^2 \\ &= \sum_{t=1}^T (y_t - \bar{y})^2 + \sum_{t=1}^T \hat{\gamma}^2 (x_t - \bar{x})^2 - 2\hat{\gamma} \sum_{t=1}^T (x_t - \bar{x})(y_t - \bar{y}) \\ T^{-1} \sum_{t=1}^T e_t^2 &= S_{yy} + \frac{S_{xy}^2}{S_{xx}^2} S_{xx} - 2 \frac{S_{xy}}{S_{xx}} \\ &= S_{yy} - \frac{S_{xy}^2}{S_{xx}} \end{aligned}$$

Therefore,

$$\begin{aligned} R^2 &= 1 - \frac{S_{yy} - \frac{S_{xy}^2}{S_{xx}}}{S_{yy}} \\ &= 1 - 1 + \frac{S_{xy}^2}{S_{xx} S_{yy}} \\ &= r_{xy}^2 \end{aligned}$$

d.)

$$\begin{aligned} F &= \frac{\hat{\gamma}^2 T S_{xx}}{s^2} = \frac{T S_{xy}^2}{S_{xx} (S_{yy} - S_{xy}^2/S_{xx}) \frac{T}{T-2}} \\ &= (T-2) \left(\frac{S_{xy}^2}{S_{xx} S_{yy} - S_{xy}^2} \right) \end{aligned}$$

e.)

$$\begin{aligned}
 F &= (T-2) \left(\frac{1}{-1 + \frac{S_{xx}S_{yy}}{S_{xy}^2}} \right) \\
 &= (T-2) \left(\frac{1}{-1 + 1/R^2} \right) \\
 &= (T-2) \left(\frac{R^2}{1 - R^2} \right)
 \end{aligned}$$

- 3.) a.) Factors in ε_i such as an economic shock might be positively correlated with s_i . This is expected to result in $\hat{\beta}_2$ being biased upward because if $cov(s_i, \varepsilon_i) > 0$

$$\hat{\beta}_2 = \beta_2 + \frac{\sum_{i=1}^n (s_i - \bar{s})(\varepsilon_i - \bar{\varepsilon})}{\sum_{i=1}^n (s_i - \bar{s})^2} \xrightarrow{p} \beta_2 + \frac{cov(s_i, \varepsilon_i)}{var(s_i)} > \beta_2$$

b.) Let $\underline{x}_i = (1, x_i)'$. $E[\underline{x}_i(1, s_i)]$ has rank 2.

c.) $E x_i \varepsilon_i = 0$

d.) Let $\hat{\underline{x}}_i = (1, \hat{s}_i)'$, $\hat{s}_i = \hat{\alpha}_1 + \hat{\alpha}_2 x_i$ and $\hat{\alpha} = (\hat{\alpha}_1, \hat{\alpha}_2)' = (\sum_{i=1}^n \hat{\underline{x}}_i \hat{\underline{x}}_i')^{-1} \sum_{i=1}^n \hat{\underline{x}}_i s_i$.

$$\hat{\beta} = \left(\sum_{i=1}^n \hat{\underline{x}}_i \hat{\underline{x}}_i' \right)^{-1} \sum_{i=1}^n \hat{\underline{x}}_i y_i$$

4.) a.) $E[y_t | y_{t-1}, \dots, y_1, x_t, \dots, x_1] = 0$ implies $E[y_t | y_{t-1}, \dots, y_1] = 0$.

b.) Score $h(\alpha, w_t) = \frac{\partial}{\partial \alpha} [-(1/2) \log(2\pi) - (1/2) \log(\alpha x_t) - y_t^2 / (2\alpha x_t)]$.

$$\begin{aligned}
 h(\alpha, w_t) &= \frac{-1}{2\alpha x_t} x_t + \frac{y_t^2}{2\alpha^2 x_t} \\
 &= \frac{-1}{2\alpha} + \frac{y_t^2}{2\alpha^2 x_t} \\
 E[h(\alpha_0, w_t)] &= \frac{-1}{2\alpha_0} + \frac{1}{2\alpha_0^2} E \left[\frac{y_t^2}{x_t} \right] \\
 &= 0
 \end{aligned}$$

c.) $\hat{\alpha} = T^{-1} \sum_{t=1}^T y_t^2 / x_t$

d.) $E[\hat{\alpha}] = \alpha$ and $var[\hat{\alpha}] = E[\hat{\alpha} - E[\hat{\alpha}]]^2 = 2\alpha^2$. Therefore,

$$\sqrt{T}(\hat{\alpha} - \alpha) \xrightarrow{L} N(0, 2\alpha^2)$$

e.) From $E[h(\alpha_0, w_t)] = 0$, GMM estimator $\hat{\alpha}$ make sample moment $T^{-1} \sum_{t=1}^T h(\hat{\alpha}, w_t) = 0$. That is,

$$\frac{1}{T} \sum_{t=1}^T \left(\frac{-1}{2\alpha} + \frac{y_t^2}{2\hat{\alpha}^2 x_t} \right) = 0$$

f.) As $E[h_t h_s] = 0$ for $t \neq s$,

$$\begin{aligned} S &= \sum_{v=-\infty}^{\infty} E[h(\alpha_0, w_t)h(\alpha_0, w_{t-v})'] \\ &= E[h(\alpha_0, w_t)h(\alpha_0, w_t)'] \\ &= \frac{1}{4\alpha^2} + \frac{3\alpha^2}{4\alpha^4} - \frac{\alpha}{2\alpha^3} \\ &= \frac{1}{2\alpha^2} \end{aligned}$$

g.)

$$\begin{aligned} D' &= E \left[\frac{\partial h(\alpha, w_t)}{\partial \alpha'} \right] = E \left[\frac{1}{2\alpha^2} - \frac{y_t^2}{\alpha^3 x_t} \right] \\ &= -\frac{1}{2\alpha^2} \\ V &= (DS^{-1}D')^{-1} \\ &= 2\alpha^2 \end{aligned}$$

h.) Yes, because the moment condition $E[h(\alpha_0, w_t)] = 0$ would still hold.

i.) $\hat{V} = [\hat{D}\hat{S}^{-1}\hat{D}]^{-1}$, where

$$\begin{aligned} \hat{D} &= -\frac{1}{2\hat{\alpha}^2} \\ \hat{S} &= \frac{1}{4\hat{\alpha}^2} + \frac{1}{4\hat{\alpha}^4} \frac{1}{T} \sum_{t=1}^T \frac{y_t^4}{x_t^2} - \frac{1}{2\hat{\alpha}^3} \frac{1}{T} \sum_{t=1}^T \frac{y_t^2}{x_t} \\ &= \frac{1}{4\hat{\alpha}^4} \frac{1}{T} \sum_{t=1}^T \frac{y_t^4}{x_t^2} - \frac{1}{4\hat{\alpha}^2} \end{aligned}$$