

No class Wed Oct 18

# Set identification using sign restrictions

Could we still draw structural conclusions using much weaker identifying assumptions, e.g., supply curve slopes up and demand curve slopes down?

$\boldsymbol{\varepsilon}_t$  = vector of VAR forecast errors

$$\boldsymbol{\Omega} = E(\boldsymbol{\varepsilon}_t \boldsymbol{\varepsilon}_t')$$

$$\hat{\boldsymbol{\varepsilon}}_t = \mathbf{y}_t - \hat{\mathbf{c}} - \hat{\boldsymbol{\Phi}}_1 \mathbf{y}_{t-1} - \cdots - \hat{\boldsymbol{\Phi}}_p \mathbf{y}_{t-p}$$

$$\hat{\boldsymbol{\Omega}} = T^{-1} \sum_{t=1}^T \hat{\boldsymbol{\varepsilon}}_t \hat{\boldsymbol{\varepsilon}}_t'$$

$\mathbf{v}_t$  = vector of structural shocks

$$E(\mathbf{v}_t \mathbf{v}_t') = \mathbf{I}_n$$

$$\boldsymbol{\varepsilon}_t = \mathbf{H} \mathbf{v}_t$$

$$E(\boldsymbol{\varepsilon}_t \boldsymbol{\varepsilon}_t') = \boldsymbol{\Omega} = \mathbf{H} \mathbf{H}'$$

$$E(\boldsymbol{\varepsilon}_t \boldsymbol{\varepsilon}_t') = \boldsymbol{\Omega} = \mathbf{H}\mathbf{H}'$$

One example of an  $\mathbf{H}$  we could consider is

Cholesky factor  $\boldsymbol{\Omega} = \mathbf{P}\mathbf{P}'$  for  $\mathbf{P}$  lower triangular.

The set of all possible  $\mathbf{H}$  can be characterized

as  $\mathbf{H} = \mathbf{P}\mathbf{Q}$  for  $\mathbf{Q} \in O_n$ , the set of all orthonormal  
( $n \times n$ ) matrices

$$O_n = \{\mathbf{Q} : \mathbf{Q}\mathbf{Q}' = \mathbf{I}_n\}$$

Proof:

(1) If  $\mathbf{H} = \mathbf{PQ}$  then  $\mathbf{HH}' = \mathbf{PQQ}'\mathbf{P}' = \mathbf{\Omega}$

(2) If  $\mathbf{HH}' = \mathbf{\Omega}$  then  $\mathbf{HH}' = \mathbf{PP}'$  and

$\mathbf{P}^{-1}\mathbf{HH}'(\mathbf{P}')^{-1} = \mathbf{I}_n$  so  $\mathbf{P}^{-1}\mathbf{H} = \mathbf{Q}$  must be an orthonormal matrix (that is,  $\mathbf{H}$  must be of form  $\mathbf{H} = \mathbf{PQ}$ )

What does  $O_n$  look like for  $n = 2$ ?

$$\mathbf{Q} = \begin{bmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{bmatrix} \text{ or } \begin{bmatrix} \cos \theta & \sin \theta \\ \sin \theta & -\cos \theta \end{bmatrix}$$

for  $\theta \in [-\pi, \pi]$ .

If we generated  $\theta \sim U[-\pi, \pi]$  and then selected one of the above matrices with prob  $1/2$ , this is described as a distribution over  $O_2$  that is Haar-uniform.

Rubio-Ramírez, Waggoner and Zha (2010)  
algorithm for generating a Haar-uniform  
draw from  $O_n$ .

(1) Generate an  $(n \times n)$  matrix  $\mathbf{X}$  of  
independent  $N(0, 1)$  variables.



(2) Calculate the  $QR$  decomposition  $\mathbf{X} = \mathbf{QR}$  where  $\mathbf{Q}$  is orthonormal and  $\mathbf{R}$  is upper triangular

Matlab:  $[\mathbf{Q}, \mathbf{R}] = \text{qr}(\mathbf{X})$

How the  $QR$  decomposition works: first column of  $\mathbf{Q}$  is simply first column of  $\mathbf{X}$  normalized to have unit length:

$$\begin{bmatrix} q_{11} \\ q_{21} \\ \vdots \\ q_{n1} \end{bmatrix} = \begin{bmatrix} x_{11}/\sqrt{x_{11}^2 + \cdots + x_{n1}^2} \\ x_{21}/\sqrt{x_{11}^2 + \cdots + x_{n1}^2} \\ \vdots \\ x_{n1}/\sqrt{x_{11}^2 + \cdots + x_{n1}^2} \end{bmatrix}$$

$$\begin{bmatrix} q_{11} \\ q_{21} \end{bmatrix} = \begin{bmatrix} x_{11}/\sqrt{x_{11}^2 + x_{21}^2} \\ x_{21}/\sqrt{x_{11}^2 + x_{21}^2} \end{bmatrix}$$

For  $n = 2$ ,  $q_{11}$  is cosine of angle formed by  $x_{11}, x_{21}$  and  $q_{21}$  is the sine.

$$\mathbf{Q} = \begin{bmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{bmatrix} \text{ or } \begin{bmatrix} \cos \theta & \sin \theta \\ \sin \theta & -\cos \theta \end{bmatrix}$$

$$\theta \sim U(-\pi, \pi)$$

Algorithm for generating possible draws for  $\mathbf{H}$ .

(1) Either fix  $\mathbf{\Omega}$  and  $\mathbf{\Gamma}$  at MLEs  $\hat{\mathbf{\Omega}} = T^{-1} \sum_{t=1}^T \hat{\mathbf{\epsilon}}_t \hat{\mathbf{\epsilon}}_t'$

and  $\hat{\mathbf{\Gamma}}' = \left( \sum_{t=1}^T \mathbf{y}_t \mathbf{x}_t' \right) \left( \sum_{t=1}^T \mathbf{x}_t \mathbf{x}_t' \right)^{-1} \Rightarrow \hat{\Psi}_s$  or

draw  $\mathbf{\Omega}^{-1}$  from Wishart with  $T - p$  degrees

of freedom and scale matrix  $T\hat{\mathbf{\Omega}}$  and use this

to draw  $\text{vec}(\mathbf{\Gamma}) \sim N\left(\text{vec}[\hat{\mathbf{\Gamma}}], \mathbf{\Omega} \otimes \left[ \sum_{t=1}^T \mathbf{x}_t \mathbf{x}_t' \right]^{-1}\right)$ .

(2) Find Cholesky factor  $\mathbf{\Omega} = \mathbf{P}\mathbf{P}'$ , draw  $\mathbf{Q}$  from Haar distribution, and calculate candidate  $\mathbf{H} = \mathbf{P}\mathbf{Q}$ .

(3) Calculate signs of chosen magnitudes in  $\Psi_s \mathbf{H}$  and keep draw if these satisfy theory, otherwise throw out.

(e.g., monetary contraction raises interest rate, lowers output and inflation on impact ( $s = 0$ ))

(4) Researchers typically report median accepted draw for element  $i, j$  of  $\Psi_s \mathbf{H}$  as if it is estimate of effect of structural shock  $j$  on variable  $i$  and 68% of draws around the median as if they were "error bands" (this is problematic!)

Example:

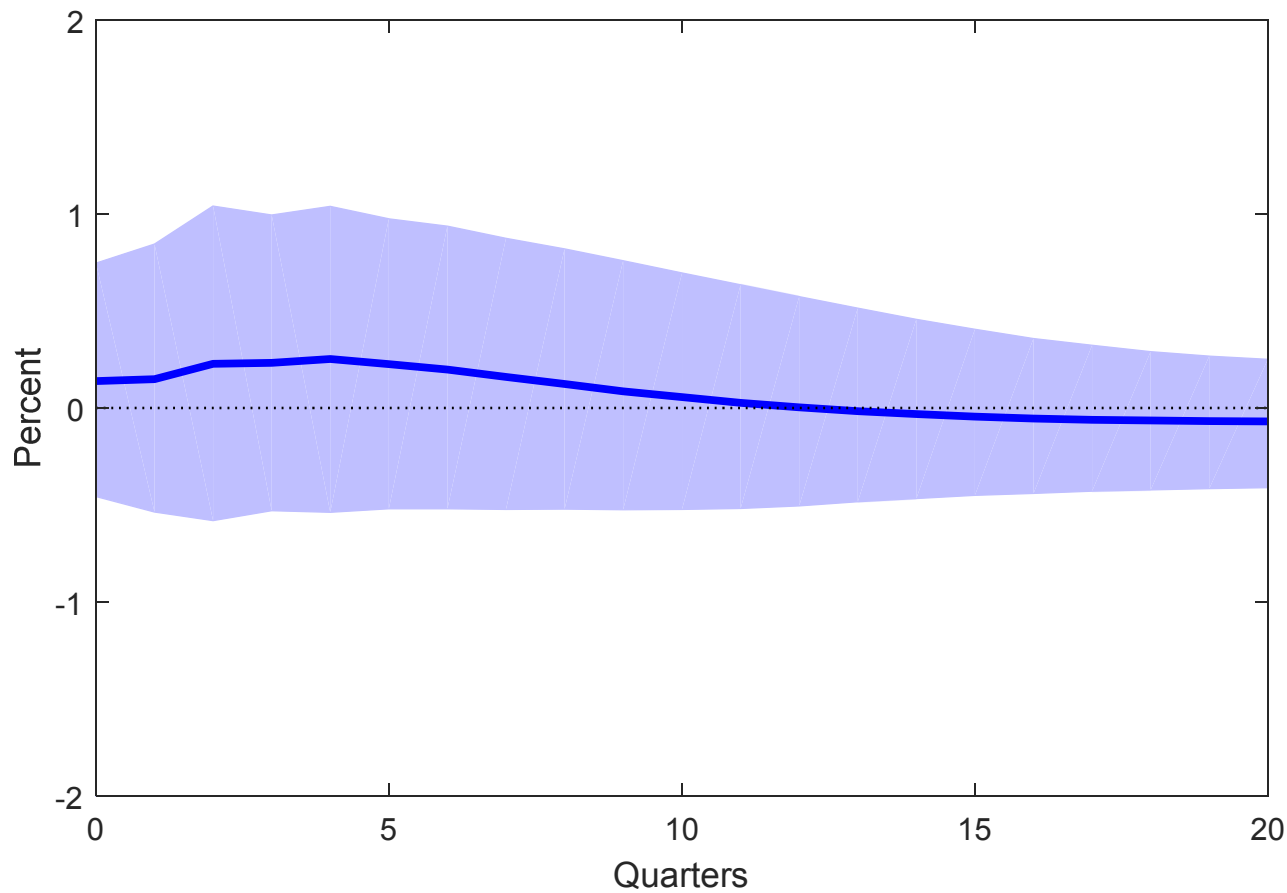
$y_{1t}$  = fed funds rate

$y_{2t}$  = log output gap

$y_{3t}$  = inflation

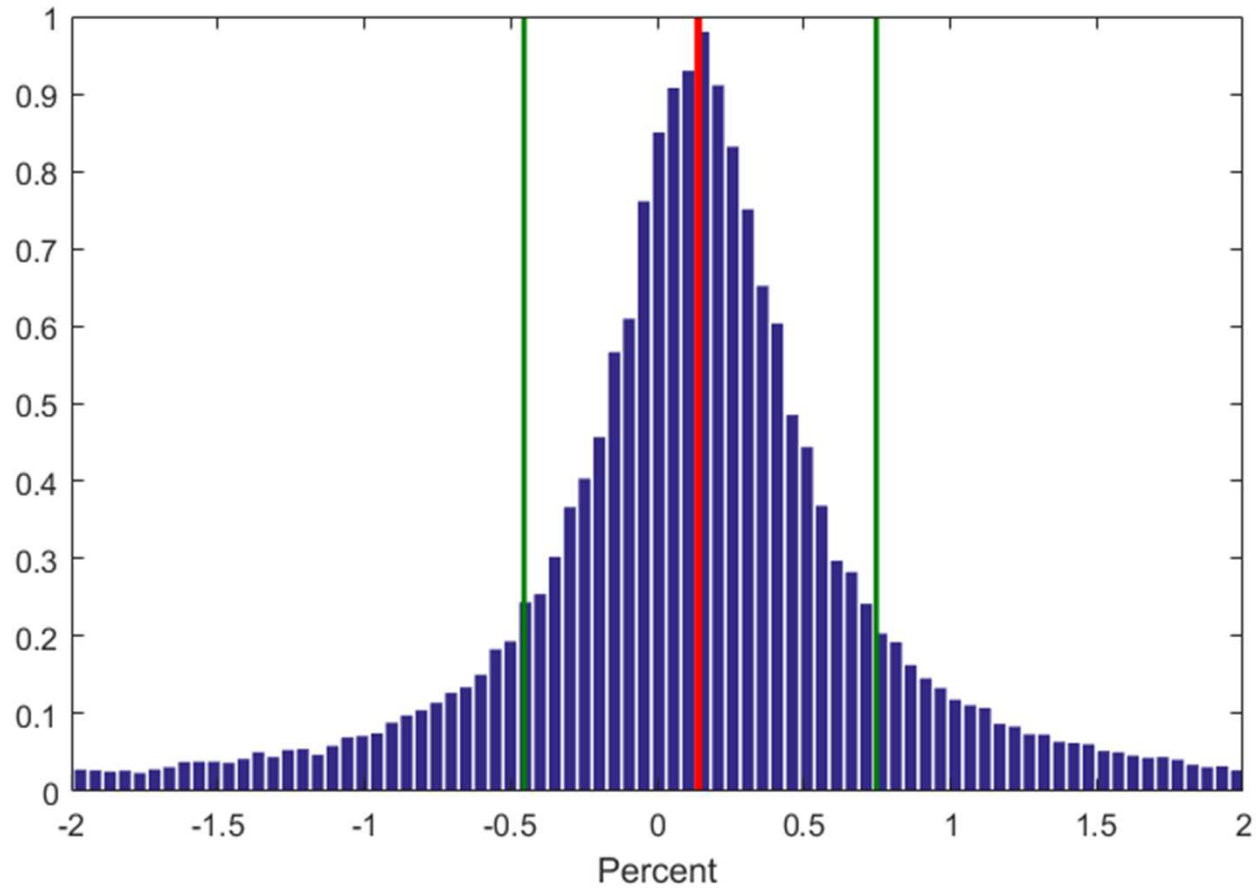
Let's run the algorithm to find the effect on output of a monetary policy shock that raises fed funds rate by 0.25%, with one change— we forget to throw any of the draws out!

Supposed effect on output gap of 0.25% monetary contraction without making any assumptions (68% “error bands”)





# Supposed effect at horizon zero on output gap of 0.25% monetary contraction



This magnitude is 0.25 times the ratio of the (2,1) element of  $\mathbf{H}$  to the (1,1) element  
= 0.25 times ratio of effect of shock 1 (monetary policy?) on output to its effect on fed funds rate

$$\mathbf{h}_1 = \mathbf{P}\mathbf{q}_1$$

$$= \begin{bmatrix} p_{11} & 0 & 0 \\ p_{21} & p_{22} & 0 \\ p_{31} & p_{32} & p_{33} \end{bmatrix} \begin{bmatrix} x_{11}/\sqrt{x_{11}^2 + x_{21}^2 + x_{31}^2} \\ x_{21}/\sqrt{x_{11}^2 + x_{21}^2 + x_{31}^2} \\ x_{31}/\sqrt{x_{11}^2 + x_{21}^2 + x_{31}^2} \end{bmatrix}$$

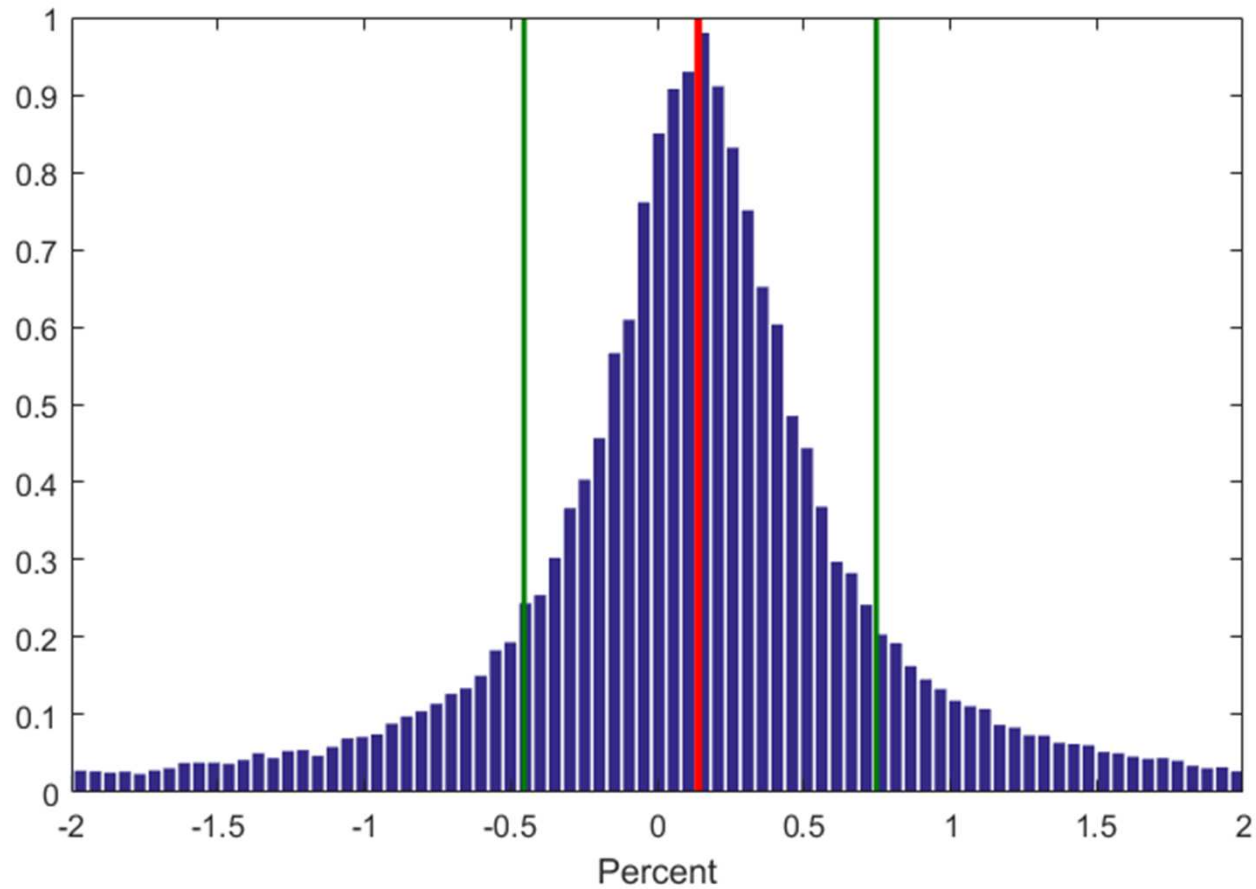
$$h_{21}/h_{11} = \frac{p_{21}x_{11} + p_{22}x_{21}}{p_{11}x_{11}} = \frac{p_{21}}{p_{11}} + \frac{p_{22}}{p_{11}} \frac{x_{21}}{x_{11}}$$

$$x_{ij} \sim N(0, 1)$$

$$x_{21}/x_{11} \sim \text{Cauchy}(0, 1)$$

$$h_{21}/h_{11} \sim \text{Cauchy}(p_{21}/p_{11}, p_{22}/p_{11})$$

# Supposed effect at horizon zero on output gap of 0.25% monetary contraction



If we reported all the draws instead of 68% "error bands," answer would just be the real line.

Implicit distribution has made it appear we learned more than we did.

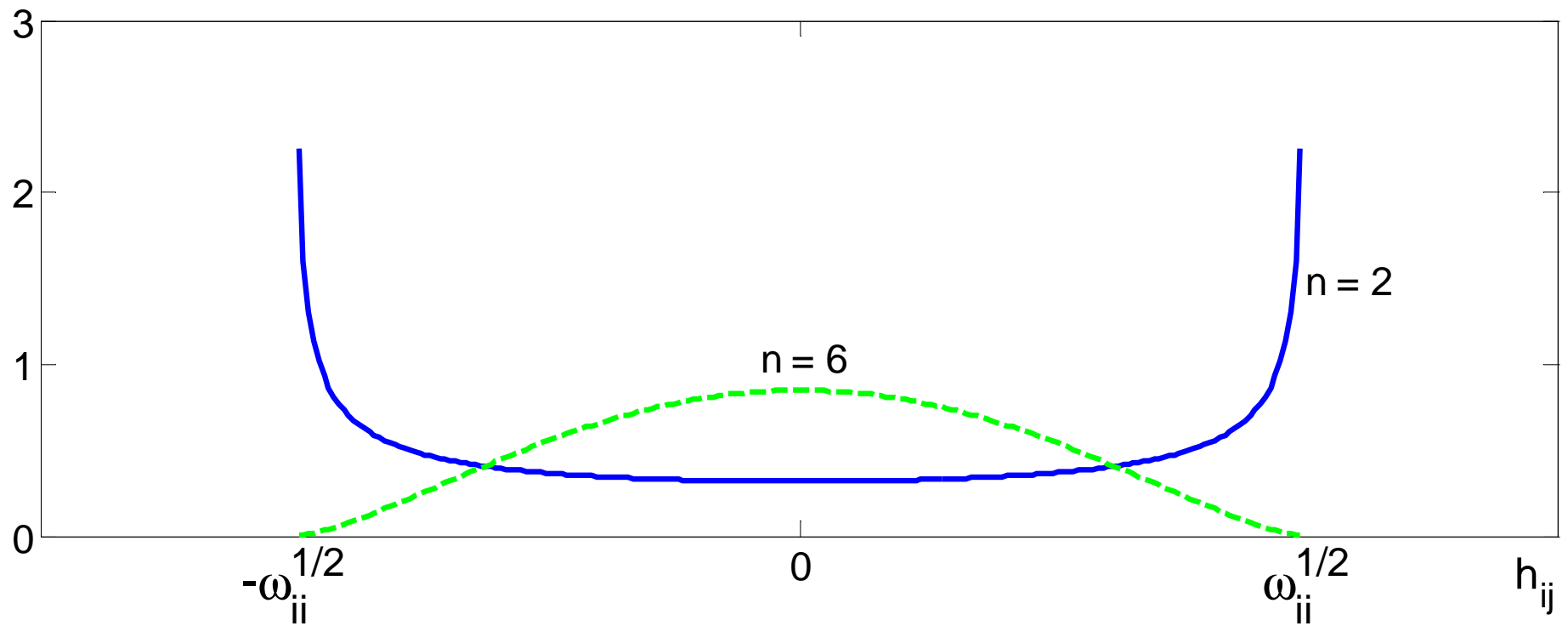
What about distribution of individual elements  $h_{ij}$ ?

$$h_{11} = p_{11}x_{11}/\sqrt{x_{11}^2 + x_{21}^2 + \cdots + x_{n1}^2}$$

$$p_{11} = \sqrt{\omega_{11}}$$

# Analytic distribution of $h_{ij}$

Panel A



Although the procedure implies a uniform distribution for the angle of rotation  $\theta$  associated with the matrix  $\mathbf{Q}$ , we are not interested in inference about  $\theta$ .

The algorithm implies a nonuniform distribution for structural impulse-response coefficients and this is what we are looking at with median and "error bands".



How do sign restrictions change any of this?

$\Delta w_t$  = growth rate of real labor compensation

$\Delta n_t$  = growth rate of total employment

$$\mathbf{y}_t = (\Delta w_t, \Delta n_t)'$$

$$\begin{aligned} \text{demand: } \Delta n_t = & k^d + \beta^d \Delta w_t + b_{11}^d \Delta w_{t-1} + b_{12}^d \Delta n_{t-1} + b_{21}^d \Delta w_{t-2} \\ & + b_{22}^d \Delta n_{t-2} + \cdots + b_{m1}^d \Delta w_{t-m} + b_{m2}^d \Delta n_{t-m} + u_t^d \end{aligned}$$

supply:

$$\begin{aligned} \Delta n_t = & k^s + \alpha^s \Delta w_t + b_{11}^s \Delta w_{t-1} + b_{12}^s \Delta n_{t-1} + b_{21}^s \Delta w_{t-2} \\ & + b_{22}^s \Delta n_{t-2} + \cdots + b_{m1}^s \Delta w_{t-m} + b_{m2}^s \Delta n_{t-m} + u_t^s \end{aligned}$$

sign restrictions:  $\beta^d \leq 0$ ,  $\alpha^s \geq 0$ .

For fixed  $\alpha^s$ , MLE of  $\beta^d$  can be found by an IV regression of  $\hat{\varepsilon}_{2t}$  on  $\hat{\varepsilon}_{1t}$  using  $\hat{\varepsilon}_{2t} - \alpha\hat{\varepsilon}_{1t}$  as instrument:

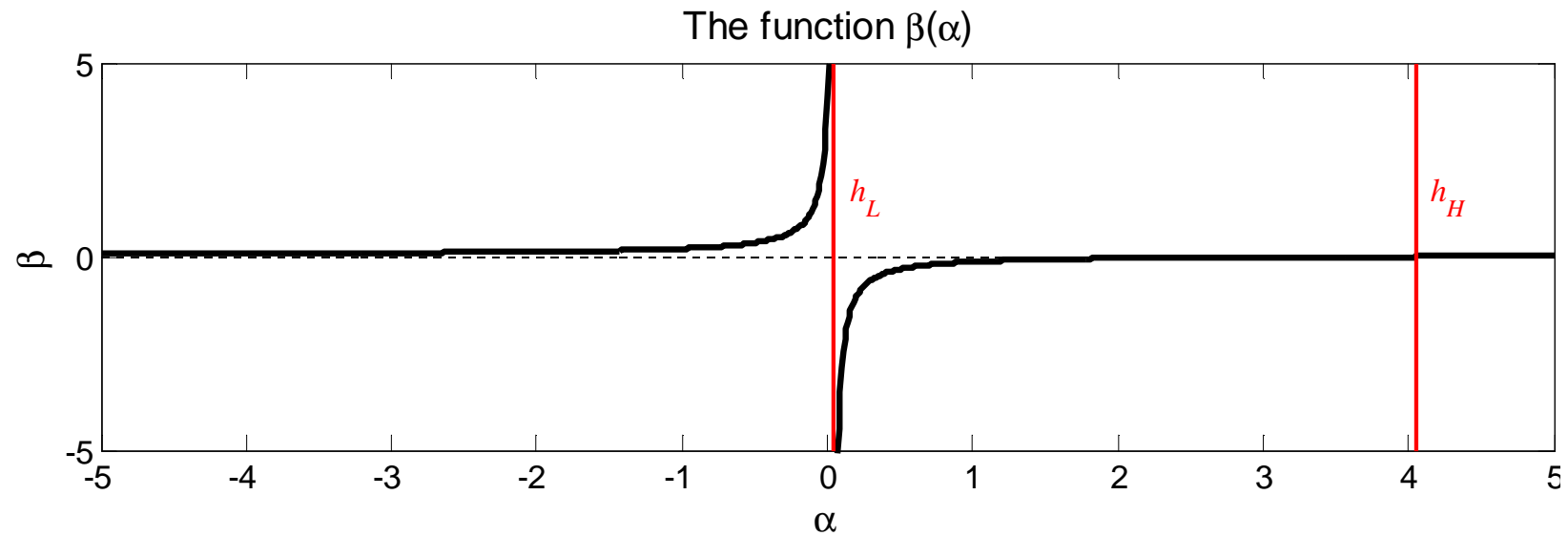
$$\hat{\beta}(\alpha) = \frac{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha\hat{\varepsilon}_{1t})\hat{\varepsilon}_{2t}}{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha\hat{\varepsilon}_{1t})\hat{\varepsilon}_{1t}} = \frac{(\hat{\omega}_{22} - \alpha\hat{\omega}_{12})}{(\hat{\omega}_{12} - \alpha\hat{\omega}_{11})}$$

$$\hat{\beta}(\alpha) = \frac{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha \hat{\varepsilon}_{1t}) \hat{\varepsilon}_{2t}}{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha \hat{\varepsilon}_{1t}) \hat{\varepsilon}_{1t}} = \frac{(\hat{\omega}_{22} - \alpha \hat{\omega}_{12})}{(\hat{\omega}_{12} - \alpha \hat{\omega}_{11})}$$

In the data,  $\hat{\omega}_{12} > 0$ .

At  $\alpha = h_H = \hat{\omega}_{22}/\hat{\omega}_{12}$ , numerator switches from positive to negative.

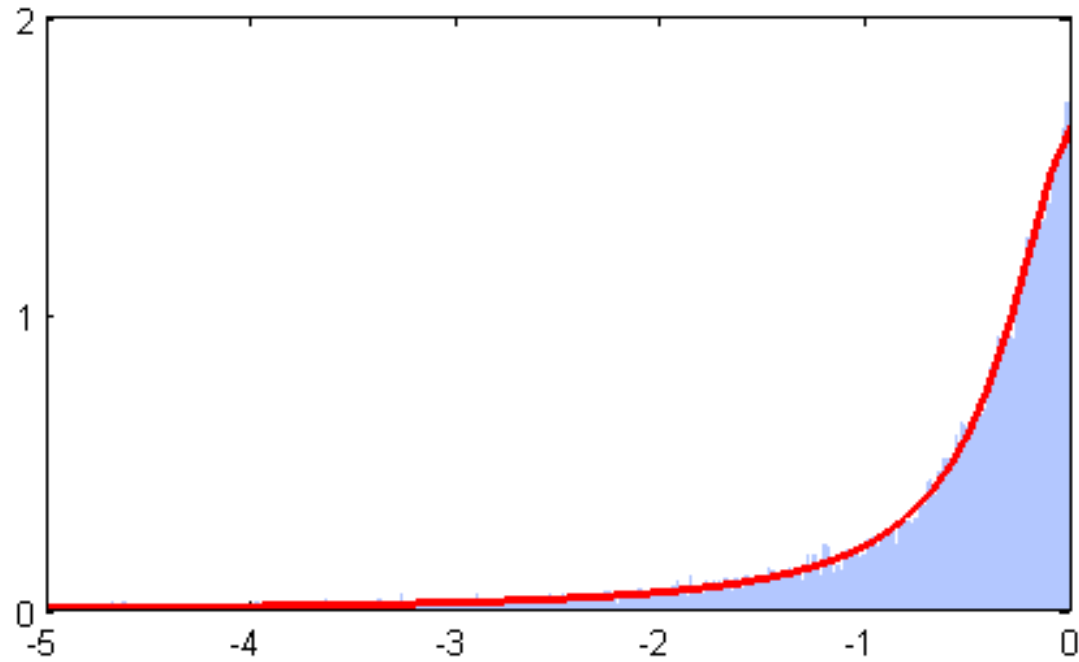
At  $\alpha = h_L = \hat{\omega}_{12}/\hat{\omega}_{11}$ , denominator switches from positive to negative.



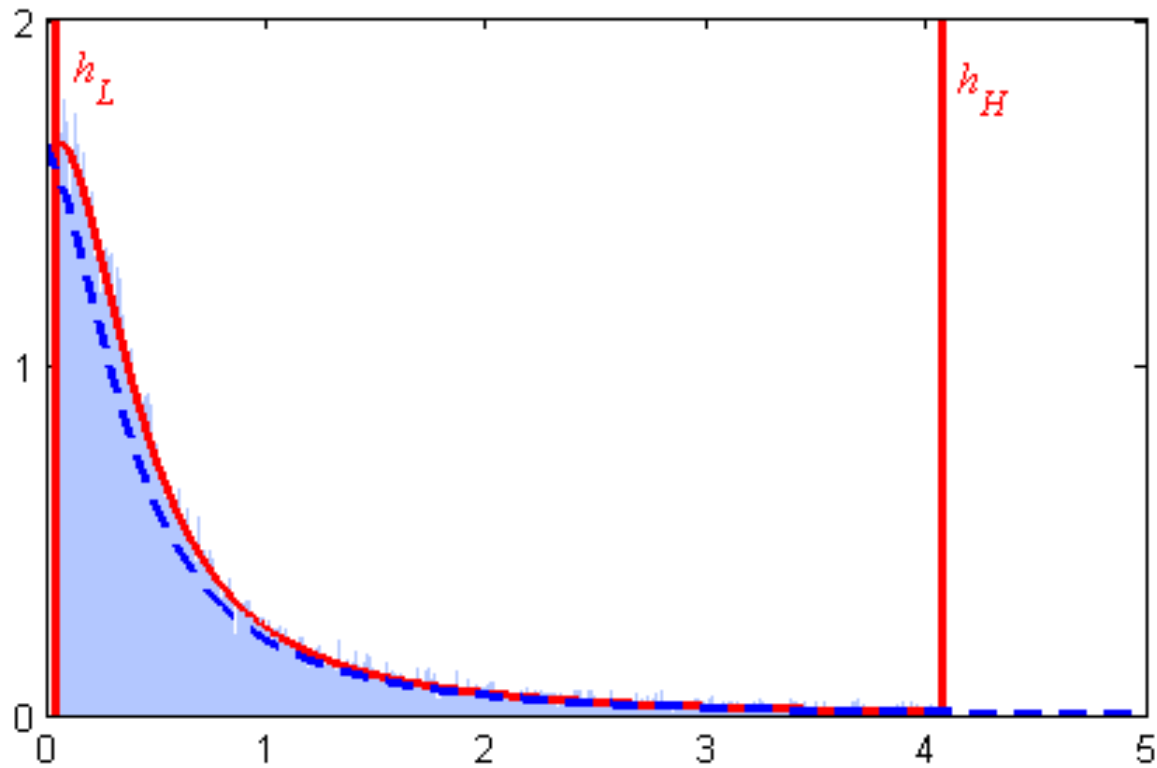
$\alpha > 0$  and  $\beta < 0$  restricts  $h_L < \alpha < h_H$   
but allows any  $\beta < 0$ .

Intuition:  $h_L = \hat{\omega}_{12}/\hat{\omega}_{11}$  is coeff from OLS  
regression of  $\hat{\varepsilon}_{2t}$  on  $\hat{\varepsilon}_{1t}$   
= convex combination of  $\alpha$  and  $\beta$   
 $\Rightarrow \beta < h_L, \alpha > h_L$   
since  $h_L > 0$ , this restricts  $\alpha$ , not  $\beta$

Intuition:  $h_H^{-1} = \hat{\omega}_{12}/\hat{\omega}_{22}$  is coefficient  
 from OLS regression of  $\hat{\varepsilon}_{1t}$  on  $\hat{\varepsilon}_{2t}$   
 = convex combination of  $\alpha^{-1}$  and  $\beta^{-1}$   
 $\Rightarrow \beta^{-1} < h_H^{-1}, \alpha^{-1} > h_H^{-1}$   
 since  $h_H > 0$ , this restricts  $\alpha$ , not  $\beta$   
 $\Rightarrow h_L < \alpha < h_H$



Distribution for draws of  $\beta$  when sign restrictions are imposed is Cauchy truncated to be negative.



Distribution for draws of  $\alpha$  when sign restrictions are imposed is Cauchy truncated to be between  $h_L$  and  $h_H$ .