

No class Wed Oct 18

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Set identification using sign restrictions

Could we still draw structural conclusions using much weaker identifying assumptions, e.g., supply curve slopes up and demand curve slopes down?

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ε_t = vector of VAR forecast errors

$$\Omega = E(\varepsilon_t \varepsilon_t')$$

$$\hat{\varepsilon}_t = \mathbf{y}_t - \hat{\mathbf{c}} - \hat{\Phi}_1 \mathbf{y}_{t-1} - \dots - \hat{\Phi}_p \mathbf{y}_{t-p}$$

$$\hat{\Omega} = T^{-1} \sum_{t=1}^T \hat{\varepsilon}_t \hat{\varepsilon}_t'$$

\mathbf{v}_t = vector of structural shocks

$$E(\mathbf{v}_t \mathbf{v}_t') = \mathbf{I}_n$$

$$\varepsilon_t = \mathbf{H} \mathbf{v}_t$$

$$E(\varepsilon_t \varepsilon_t') = \Omega = \mathbf{H} \mathbf{H}'$$

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$$E(\varepsilon_t \varepsilon_t') = \Omega = \mathbf{H} \mathbf{H}'$$

One example of an \mathbf{H} we could consider is Cholesky factor $\Omega = \mathbf{P} \mathbf{P}'$ for \mathbf{P} lower triangular.

The set of all possible \mathbf{H} can be characterized as $\mathbf{H} = \mathbf{P} \mathbf{Q}$ for $\mathbf{Q} \in O_n$, the set of all orthonormal ($n \times n$) matrices

$$O_n = \{\mathbf{Q} : \mathbf{Q} \mathbf{Q}' = \mathbf{I}_n\}$$

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Proof:

(1) If $\mathbf{H} = \mathbf{P} \mathbf{Q}$ then $\mathbf{H} \mathbf{H}' = \mathbf{P} \mathbf{Q} \mathbf{Q}' \mathbf{P}' = \Omega$

(2) If $\mathbf{H} \mathbf{H}' = \Omega$ then $\mathbf{H} \mathbf{H}' = \mathbf{P} \mathbf{P}'$ and $\mathbf{P}^{-1} \mathbf{H} \mathbf{H}' (\mathbf{P}')^{-1} = \mathbf{I}_n$ so $\mathbf{P}^{-1} \mathbf{H} = \mathbf{Q}$ must be an orthonormal matrix (that is, \mathbf{H} must be of form $\mathbf{H} = \mathbf{P} \mathbf{Q}$)

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What does O_n look like for $n = 2$?

$$\mathbf{Q} = \begin{bmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{bmatrix} \text{ or } \begin{bmatrix} \cos \theta & \sin \theta \\ \sin \theta & -\cos \theta \end{bmatrix}$$

for $\theta \in [-\pi, \pi]$.

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If we generated $\theta \sim U[-\pi, \pi]$ and then selected one of the above matrices with prob 1/2, this is described as a distribution over O_2 that is Haar-uniform.

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Rubio-Ramírez, Waggoner and Zha (2010) algorithm for generating a Haar-uniform draw from O_n .

(1) Generate an $(n \times n)$ matrix \mathbf{X} of independent $N(0, 1)$ variables.

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(2) Calculate the QR decomposition $\mathbf{X} = \mathbf{QR}$ where \mathbf{Q} is orthonormal and \mathbf{R} is upper triangular
Matlab: `[Q,R] = qr(X)`

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How the QR decomposition works: first column of \mathbf{Q} is simply first column of \mathbf{X} normalized to have unit length:

$$\begin{bmatrix} q_{11} \\ q_{21} \\ \vdots \\ q_{n1} \end{bmatrix} = \begin{bmatrix} x_{11}/\sqrt{x_{11}^2 + \dots + x_{n1}^2} \\ x_{21}/\sqrt{x_{11}^2 + \dots + x_{n1}^2} \\ \vdots \\ x_{n1}/\sqrt{x_{11}^2 + \dots + x_{n1}^2} \end{bmatrix}$$

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$$\begin{bmatrix} q_{11} \\ q_{21} \end{bmatrix} = \begin{bmatrix} x_{11}/\sqrt{x_{11}^2 + x_{21}^2} \\ x_{21}/\sqrt{x_{11}^2 + x_{21}^2} \end{bmatrix}$$

For $n = 2$, q_{11} is cosine of angle formed by x_{11}, x_{21} and q_{21} is the sine.

$$\mathbf{Q} = \begin{bmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{bmatrix} \text{ or } \begin{bmatrix} \cos \theta & \sin \theta \\ \sin \theta & -\cos \theta \end{bmatrix}$$

$\theta \sim U(-\pi, \pi)$

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Algorithm for generating possible draws for \mathbf{H} .

(1) Either fix $\mathbf{\Omega}$ and $\mathbf{\Gamma}$ at MLEs $\hat{\mathbf{\Omega}} = T^{-1} \sum_{t=1}^T \hat{\mathbf{\epsilon}}_t \hat{\mathbf{\epsilon}}_t'$ and $\hat{\mathbf{\Gamma}}' = (\sum_{t=1}^T \mathbf{y}_t \mathbf{x}_t') (\sum_{t=1}^T \mathbf{x}_t \mathbf{x}_t')^{-1} \Rightarrow \hat{\mathbf{\Psi}}_s$ or draw $\mathbf{\Omega}^{-1}$ from Wishart with $T-p$ degrees of freedom and scale matrix $T\hat{\mathbf{\Omega}}$ and use this to draw $\text{vec}(\mathbf{\Gamma}) \sim N(\text{vec}[\hat{\mathbf{\Gamma}}], \mathbf{\Omega} \otimes [\sum_{t=1}^T \mathbf{x}_t \mathbf{x}_t']^{-1})$.

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(2) Find Cholesky factor $\Omega = \mathbf{P}\mathbf{P}'$, draw \mathbf{Q} from Haar distribution, and calculate candidate $\mathbf{H} = \mathbf{P}\mathbf{Q}$.

(3) Calculate signs of chosen magnitudes in $\Psi_s\mathbf{H}$ and keep draw if these satisfy theory, otherwise throw out.
(e.g., monetary contraction raises interest rate, lowers output and inflation on impact ($s = 0$))

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(4) Researchers typically report median accepted draw for element i, j of $\Psi_s\mathbf{H}$ as if it is estimate of effect of structural shock j on variable i and 68% of draws around the median as if they were "error bands" (this is problematic!)

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Example:

y_{1t} = fed funds rate

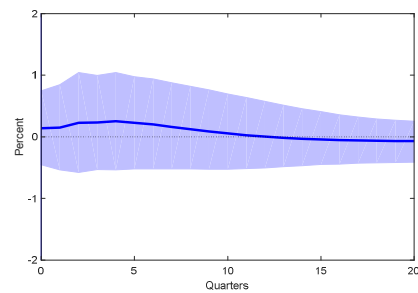
y_{2t} = log output gap

y_{3t} = inflation

Let's run the algorithm to find the effect on output of a monetary policy shock that raises fed funds rate by 0.25%, with one change—we forget to throw any of the draws out!

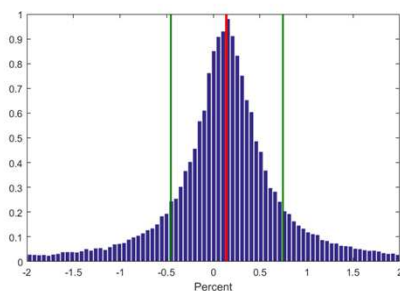
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Supposed effect on output gap of 0.25% monetary contraction without making any assumptions (68% "error bands")



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Supposed effect at horizon zero on output gap of 0.25% monetary contraction



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This magnitude is 0.25 times the ratio of the (2,1) element of \mathbf{H} to the (1,1) element = 0.25 times ratio of effect of shock 1 (monetary policy?) on output to its effect on fed funds rate

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$\mathbf{h}_1 = \mathbf{P}\mathbf{q}_1$

$$= \begin{bmatrix} p_{11} & 0 & 0 \\ p_{21} & p_{22} & 0 \\ p_{31} & p_{32} & p_{33} \end{bmatrix} \begin{bmatrix} x_{11}/\sqrt{x_{11}^2 + x_{21}^2 + x_{31}^2} \\ x_{21}/\sqrt{x_{11}^2 + x_{21}^2 + x_{31}^2} \\ x_{31}/\sqrt{x_{11}^2 + x_{21}^2 + x_{31}^2} \end{bmatrix}$$

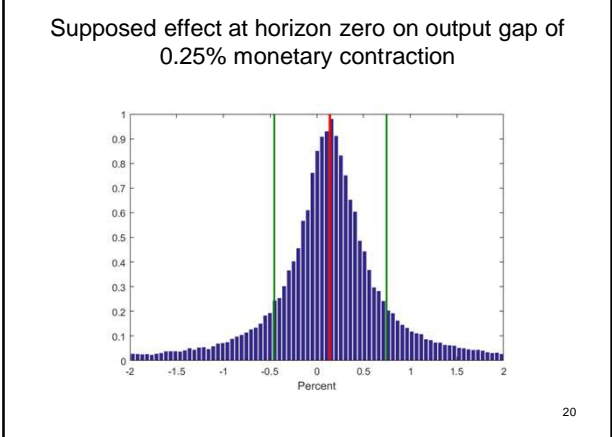
$h_{21}/h_{11} = \frac{p_{21}x_{11} + p_{22}x_{21}}{p_{11}x_{11}} = \frac{p_{21}}{p_{11}} + \frac{p_{22}}{p_{11}} \frac{x_{21}}{x_{11}}$

$x_{ij} \sim N(0, 1)$

$x_{21}/x_{11} \sim \text{Cauchy}(0, 1)$

$h_{21}/h_{11} \sim \text{Cauchy}(p_{21}/p_{11}, p_{22}/p_{11})$

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If we reported all the draws instead of 68% "error bands," answer would just be the real line.

Implicit distribution has made it appear we learned more than we did.

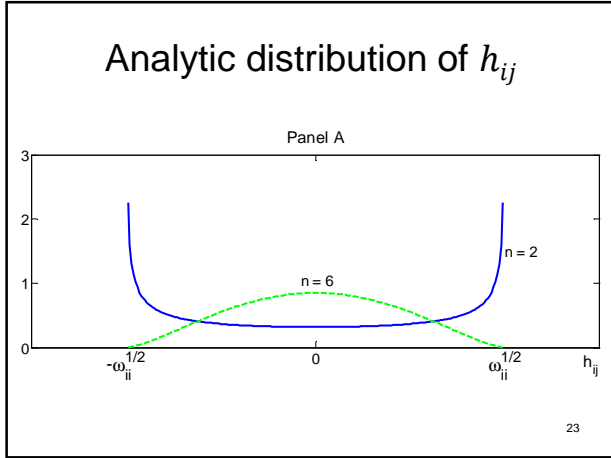
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What about distribution of individual elements h_{ij} ?

$$h_{11} = p_{11}x_{11}/\sqrt{x_{11}^2 + x_{21}^2 + \dots + x_{n1}^2}$$

$$p_{11} = \sqrt{\omega_{11}}$$

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Although the procedure implies a uniform distribution for the angle of rotation θ associated with the matrix \mathbf{Q} , we are not interested in inference about θ .

The algorithm implies a nonuniform distribution for structural impulse-response coefficients and this is what we are looking at with median and "error bands".

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How do sign restrictions change any of this?

Δw_t = growth rate of real labor compensation

Δn_t = growth rate of total employment

$y_t = (\Delta w_t, \Delta n_t)'$

demand: $\Delta n_t = k^d + \beta^d \Delta w_t + b_{11}^d \Delta w_{t-1} + b_{12}^d \Delta n_{t-1} + b_{21}^d \Delta w_{t-2}$
 $+ b_{22}^d \Delta n_{t-2} + \dots + b_{m1}^d \Delta w_{t-m} + b_{m2}^d \Delta n_{t-m} + u_t^d$

supply:

$\Delta n_t = k^s + \alpha^s \Delta w_t + b_{11}^s \Delta w_{t-1} + b_{12}^s \Delta n_{t-1} + b_{21}^s \Delta w_{t-2}$
 $+ b_{22}^s \Delta n_{t-2} + \dots + b_{m1}^s \Delta w_{t-m} + b_{m2}^s \Delta n_{t-m} + u_t^s$

sign restrictions: $\beta^d \leq 0, \alpha^s \geq 0$.

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For fixed α^s , MLE of β^d can be

found by an IV regression of $\hat{\varepsilon}_{2t}$

on $\hat{\varepsilon}_{1t}$ using $\hat{\varepsilon}_{2t} - \alpha \hat{\varepsilon}_{1t}$ as instrument:

$$\hat{\beta}(\alpha) = \frac{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha \hat{\varepsilon}_{1t}) \hat{\varepsilon}_{2t}}{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha \hat{\varepsilon}_{1t}) \hat{\varepsilon}_{1t}} = \frac{(\hat{\omega}_{22} - \alpha \hat{\omega}_{12})}{(\hat{\omega}_{12} - \alpha \hat{\omega}_{11})}$$

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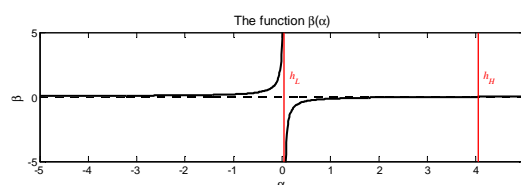
$$\hat{\beta}(\alpha) = \frac{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha \hat{\varepsilon}_{1t}) \hat{\varepsilon}_{2t}}{\sum_{t=1}^T (\hat{\varepsilon}_{2t} - \alpha \hat{\varepsilon}_{1t}) \hat{\varepsilon}_{1t}} = \frac{(\hat{\omega}_{22} - \alpha \hat{\omega}_{12})}{(\hat{\omega}_{12} - \alpha \hat{\omega}_{11})}$$

In the data, $\hat{\omega}_{12} > 0$.

At $\alpha = h_H = \hat{\omega}_{22}/\hat{\omega}_{12}$, numerator switches from positive to negative.

At $\alpha = h_L = \hat{\omega}_{12}/\hat{\omega}_{11}$, denominator switches from positive to negative.

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$\alpha > 0$ and $\beta < 0$ restricts $h_L < \alpha < h_H$ but allows any $\beta < 0$.

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Intuition: $h_L = \hat{\omega}_{12}/\hat{\omega}_{11}$ is coeff from OLS regression of $\hat{\varepsilon}_{2t}$ on $\hat{\varepsilon}_{1t}$

= convex combination of α and β

$\Rightarrow \beta < h_L, \alpha > h_L$

since $h_L > 0$, this restricts α , not β

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Intuition: $h_H^{-1} = \hat{\omega}_{12}/\hat{\omega}_{22}$ is coefficient from OLS regression of $\hat{\varepsilon}_{1t}$ on $\hat{\varepsilon}_{2t}$

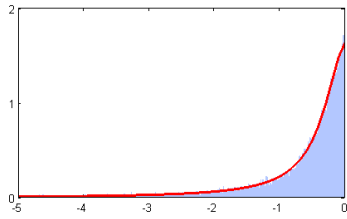
= convex combination of α^{-1} and β^{-1}

$\Rightarrow \beta^{-1} < h_H^{-1}, \alpha^{-1} > h_H^{-1}$

since $h_H > 0$, this restricts α , not β

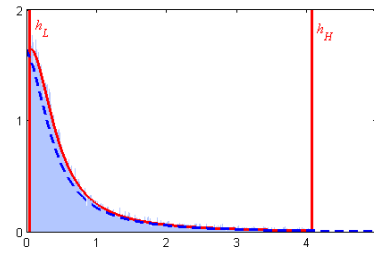
$\Rightarrow h_L < \alpha < h_H$

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Distribution for draws of β when sign restrictions are imposed is Cauchy truncated to be negative.

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Distribution for draws of α when sign restrictions are imposed is Cauchy truncated to be between h_L and h_H .

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