

New meeting times for Econ 210D

Mondays 8:00-9:20 a.m. in Econ 300

Wednesdays 11:00-12:20 in Econ 300

Identification using nonrecursive structure, long-run restrictions and heteroskedasticity

General statement of problem of structural interpretation:

Can observe in the data:

$\varepsilon_{1t}, \dots, \varepsilon_{nt}$ = errors I make forecasting variables from lagged values.

Think of these as resulting from n structural shocks:

u_{1t} = shock to technology

u_{2t} = shock to price markup

u_{3t} = shock to monetary policy

$$\boldsymbol{\varepsilon}_t = \mathbf{H}\mathbf{u}_t$$

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$$E(\boldsymbol{\varepsilon}_t \boldsymbol{\varepsilon}_t') = \boldsymbol{\Omega} \text{ (can observe in the data)}$$

$$E(\mathbf{u}_t \mathbf{u}_t') = \mathbf{D} \text{ (unknown variance of structural shocks)}$$

$$\boldsymbol{\Omega} = \mathbf{H}\mathbf{D}\mathbf{H}'$$

A. Nonrecursive Orthogonalized VARs

Structural model:

$$\mathbf{B}_0 \mathbf{y}_t = \mathbf{B} \mathbf{x}_t + \mathbf{u}_t$$

$$\mathbf{x}_t = (1, \mathbf{y}'_{t-1}, \mathbf{y}'_{t-2}, \dots, \mathbf{y}'_{t-p})'$$

\mathbf{u}_t = vector of structural shocks

$$E(\mathbf{u}_t \mathbf{u}_t') = \mathbf{D} \text{ (diagonal)}$$

recursive identification assumed \mathbf{B}_0

was lower triangular

If $\mathbf{u}_t \sim N(\mathbf{0}, \mathbf{D})$ then log likelihood is

$$-(Tn/2) \log(2\pi) - (T/2) \log |\mathbf{B}_0^{-1} \mathbf{D} (\mathbf{B}_0^{-1})'| \\ -(1/2) \sum_{t=1}^T (\mathbf{B}_0 \mathbf{y}_t - \mathbf{B} \mathbf{x}_t)' \mathbf{D}^{-1} (\mathbf{B}_0 \mathbf{y}_t - \mathbf{B} \mathbf{x}_t)$$

If \mathbf{B} is unrestricted, MLE of \mathbf{B}_0 and \mathbf{D} are values that maximize

$$(T/2) \log|\mathbf{B}_0|^2 - (T/2) \log|\mathbf{D}| \\ -(T/2) \text{trace}\{(\mathbf{B}_0' \mathbf{D}^{-1} \mathbf{B}_0) \hat{\mathbf{\Omega}}\}$$

If model is just identified, estimates will satisfy

$$\hat{\mathbf{B}}_0^{-1} \hat{\mathbf{D}} (\hat{\mathbf{B}}_0^{-1})' = \hat{\mathbf{\Omega}}$$

B. Identification using long-run conditions

x_t log of productivity in quarter t

(log GDP minus log of hours worked)

n_t log of hours worked in nonag establishments

1948:Q1-1994:Q4

$$\mathbf{y}_t = \begin{bmatrix} \Delta x_t \\ \Delta n_t \end{bmatrix} \sim I(0)$$

VAR (reduced-form) ($p = 4$)

$$\mathbf{y}_t = \mathbf{c} + \mathbf{\Phi}_1 \mathbf{y}_{t-1} + \mathbf{\Phi}_2 \mathbf{y}_{t-2} + \cdots + \mathbf{\Phi}_p \mathbf{y}_{t-p} + \boldsymbol{\varepsilon}_t$$

$$E(\boldsymbol{\varepsilon}_t \boldsymbol{\varepsilon}_t') = \boldsymbol{\Omega}$$

Structural model:

$$\mathbf{B}_0 \mathbf{y}_t = \boldsymbol{\lambda} + \mathbf{B}_1 \mathbf{y}_{t-1} + \mathbf{B}_2 \mathbf{y}_{t-2} + \cdots + \mathbf{B}_p \mathbf{y}_{t-p} + \mathbf{u}_t$$

$$E(\mathbf{u}_t \mathbf{u}_t') = \mathbf{I}_2 \quad (\text{normalization})$$

Relation between representations:

$$\mathbf{u}_t = \mathbf{B}_0 \boldsymbol{\varepsilon}_t$$

$$\boldsymbol{\Omega} = \mathbf{B}_0^{-1} (\mathbf{B}_0^{-1})'$$

Premultiply structural model,

$$\mathbf{B}(L)\mathbf{y}_t = \boldsymbol{\lambda} + \mathbf{u}_t$$

by $\mathbf{C}(L) = \mathbf{B}(L)^{-1}$:

$$\mathbf{y}_t = \boldsymbol{\mu} + \mathbf{C}_0\mathbf{u}_t + \mathbf{C}_1\mathbf{u}_{t-1} + \mathbf{C}_2\mathbf{u}_{t-2} + \dots$$

which gives structural MA representation

$$\mathbf{u}_t = \begin{bmatrix} u_{1t} \\ u_{2t} \end{bmatrix}$$

u_{1t} technology shock

u_{2t} nontechnology shock

Assumption: only technology shocks can have a permanent effect on productivity

$$\lim_{s \rightarrow \infty} \frac{\partial x_{t+s}}{\partial u_{2t}} = 0$$

Notice

$$\frac{\partial x_{t+s}}{\partial u_{2t}} = \frac{\partial(x_{t+s} - x_{t+s-1})}{\partial u_{2t}} + \frac{\partial(x_{t+s-1} - x_{t+s-2})}{\partial u_{2t}} + \dots + \frac{\partial(x_t - x_{t-1})}{\partial u_{2t}}$$

$$\mathbf{y}_t = \begin{bmatrix} x_t - x_{t-1} \\ n_t - n_{t-1} \end{bmatrix}$$

$$\frac{\partial(x_t - x_{t-1})}{\partial u_{2t}} = \frac{\partial y_{1t}}{\partial u_{2t}}$$

$$\mathbf{y}_t = \boldsymbol{\mu} + \mathbf{C}_0 \mathbf{u}_t + \mathbf{C}_1 \mathbf{u}_{t-1} + \mathbf{C}_2 \mathbf{u}_{t-2} + \dots$$

$$\frac{\partial \mathbf{y}_{t+m}}{\partial \mathbf{u}_t} = \mathbf{C}_m$$

$$\frac{\partial x_{t+s}}{\partial u_{2t}} = \frac{\partial(x_{t+s}-x_{t+s-1})}{\partial u_{2t}} + \frac{\partial(x_{t+s-1}-x_{t+s-2})}{\partial u_{2t}} + \dots + \frac{\partial(x_t-x_{t-1})}{\partial u_{2t}}$$

is given by the row 1 column 2 element of

$$\mathbf{C}_0 + \mathbf{C}_1 + \mathbf{C}_2 + \dots + \mathbf{C}_s$$

$$\lim_{s \rightarrow \infty} \frac{\partial x_{t+s}}{\partial u_{2t}} = 0$$

requires that the following matrix is lower triangular:

$$\mathbf{C}_0 + \mathbf{C}_1 + \mathbf{C}_2 + \dots = \mathbf{C}(1)$$

Goal: find structural disturbances \mathbf{u}_t that are a linear combination of the VAR innovations,

$$\mathbf{u}_t = \mathbf{B}_0 \boldsymbol{\varepsilon}_t,$$

such that:

$$(1) \ E(\mathbf{u}_t \mathbf{u}_t') = \mathbf{I}_2$$

$$\Rightarrow \mathbf{B}_0 \boldsymbol{\Omega} \mathbf{B}_0' = \mathbf{I}_2$$

$$\Rightarrow \boldsymbol{\Omega} = (\mathbf{B}_0^{-1})(\mathbf{B}_0^{-1})'$$

(2) $\mathbf{y}_t = \boldsymbol{\mu} + \mathbf{C}(L)\mathbf{u}_t$

(3) $\mathbf{C}(1)$ is lower triangular

$$\Phi(L)\mathbf{y}_t = \mathbf{c} + \boldsymbol{\varepsilon}_t$$

$$\boldsymbol{\varepsilon}_t = \mathbf{B}_0^{-1}\mathbf{u}_t$$

$$\Rightarrow \Phi(L)\mathbf{y}_t = \mathbf{c} + \mathbf{B}_0^{-1}\mathbf{u}_t$$

$$\Rightarrow \mathbf{y}_t = \boldsymbol{\mu} + [\Phi(L)]^{-1}\mathbf{B}_0^{-1}\mathbf{u}_t$$

$$\mathbf{y}_t = \boldsymbol{\mu} + \mathbf{C}(L)\mathbf{u}_t$$

$$\Rightarrow \mathbf{C}(1) = [\Phi(1)]^{-1}\mathbf{B}_0^{-1}$$

$$\mathbf{C}(1) = [\mathbf{\Phi}(1)]^{-1} \mathbf{B}_0^{-1}$$

$$\mathbf{C}(1)[\mathbf{C}(1)]' =$$

$$[\mathbf{\Phi}(1)]^{-1} \mathbf{B}_0^{-1} (\mathbf{B}_0^{-1})' \{[\mathbf{\Phi}(1)]^{-1}\}'$$

$$\mathbf{C}(1)[\mathbf{C}(1)]' = [\mathbf{\Phi}(1)]^{-1} \mathbf{\Omega} \{[\mathbf{\Phi}(1)]^{-1}\}'$$

Can estimate $\mathbf{\Phi}(1)$ and $\mathbf{\Omega}$ from VAR

$$\hat{\mathbf{\Omega}} = T^{-1} \sum_{t=1}^T \hat{\mathbf{\varepsilon}}_t \hat{\mathbf{\varepsilon}}_t'$$

$$\hat{\mathbf{\Phi}}(1) = \mathbf{I}_2 - \hat{\mathbf{\Phi}}_1 - \hat{\mathbf{\Phi}}_2 - \hat{\mathbf{\Phi}}_3 - \hat{\mathbf{\Phi}}_4$$

Want: Lower triangular matrix $\mathbf{C}(1)$ such that

$$\mathbf{C}(1)[\mathbf{C}(1)]' = [\Phi(1)]^{-1}\mathbf{\Omega}\{[\Phi(1)]^{-1}\}'$$

Conclusion: $\mathbf{C}(1)$ is Cholesky factor of

$$[\Phi(1)]^{-1}\mathbf{\Omega}\{[\Phi(1)]^{-1}\}'$$

To get \mathbf{B}_0 we then use fact that

$$\mathbf{C}(1) = [\Phi(1)]^{-1} \mathbf{B}_0^{-1}$$

$$\mathbf{B}_0 = [\mathbf{C}(1)]^{-1} [\Phi(1)]^{-1}$$

Summary:

(1) Estimate VAR by OLS

$$\mathbf{y}_t = \begin{bmatrix} \Delta x_t \\ \Delta n_t \end{bmatrix}$$

$$\mathbf{y}_t = \hat{\mathbf{c}} + \hat{\mathbf{\Phi}}_1 \mathbf{y}_{t-1} + \hat{\mathbf{\Phi}}_2 \mathbf{y}_{t-2} + \cdots + \hat{\mathbf{\Phi}}_p \mathbf{y}_{t-p} + \hat{\boldsymbol{\varepsilon}}_t$$

$$\hat{\boldsymbol{\Omega}} = T^{-1} \sum_{t=1}^T \hat{\boldsymbol{\varepsilon}}_t \hat{\boldsymbol{\varepsilon}}_t'$$

(2) Find Cholesky factor or lower triangular matrix $\hat{\mathbf{C}}$ such that

$$\hat{\mathbf{C}}\hat{\mathbf{C}}' = \hat{\mathbf{Q}}\hat{\mathbf{\Omega}}\hat{\mathbf{Q}}'$$

$$\hat{\mathbf{Q}} = \left(\mathbf{I}_2 - \hat{\mathbf{\Phi}}_1 - \hat{\mathbf{\Phi}}_2 - \dots - \hat{\mathbf{\Phi}}_p \right)^{-1}$$

(3) Technology shock and transitory shocks for date t are first and second elements of

$$\hat{\mathbf{u}}_t = \hat{\mathbf{B}}_0 \hat{\boldsymbol{\varepsilon}}_t$$

where

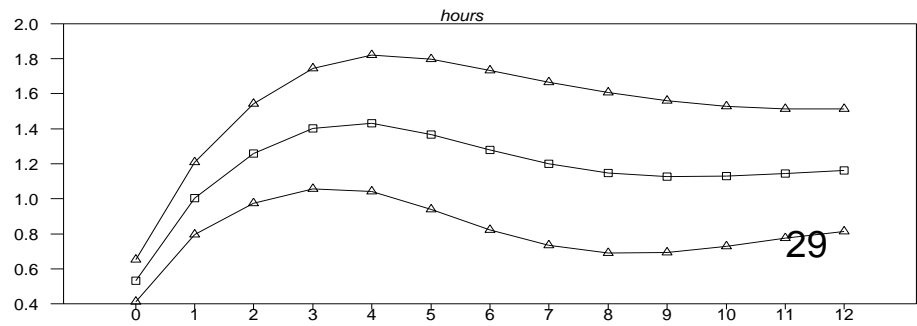
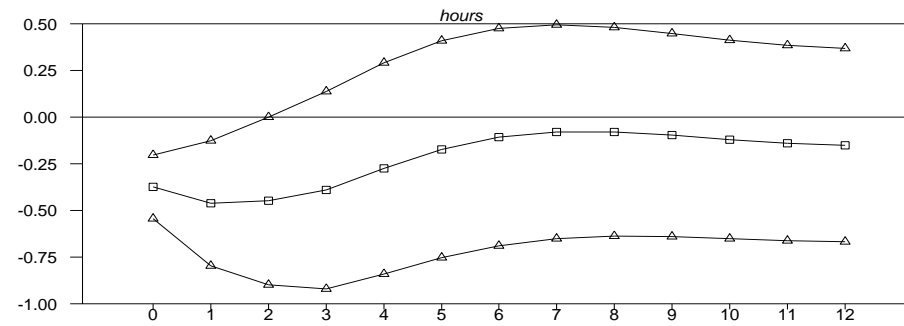
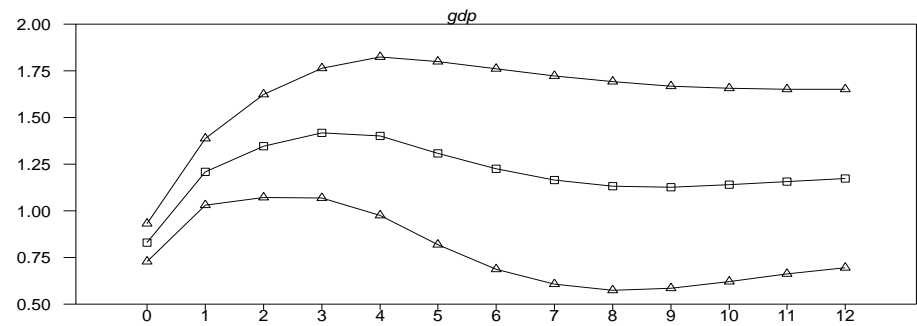
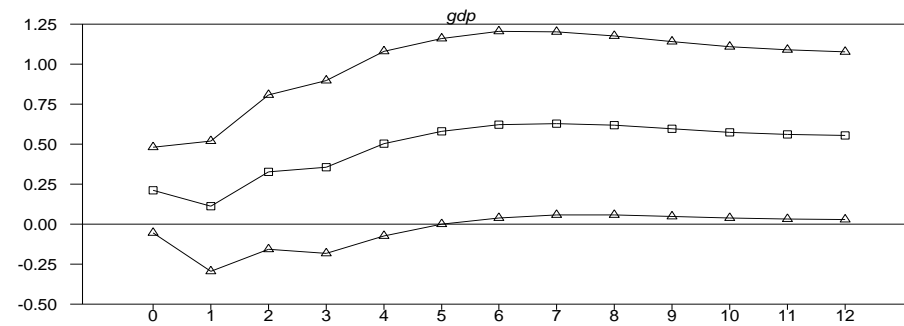
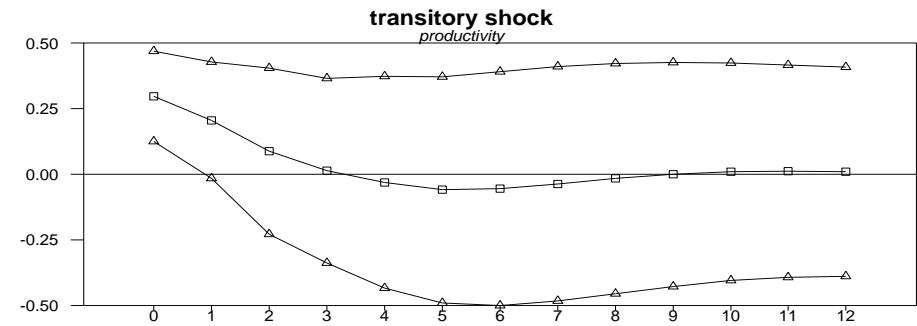
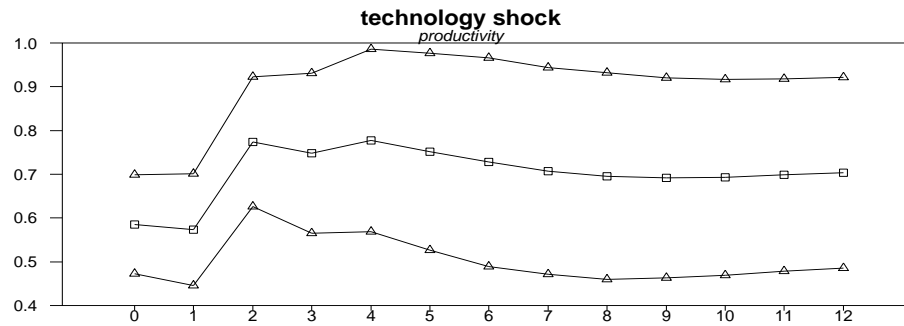
$$\hat{\mathbf{B}}_0 = \hat{\mathbf{C}}^{-1} \hat{\mathbf{Q}}$$

(4) Effect of tech shock and transitory shock at date t on \mathbf{y}_{t+s} are given by first and second columns, respectively, of

$$\frac{\partial \mathbf{y}_{t+s}}{\partial \mathbf{u}_t'} = \mathbf{\Psi}_s \mathbf{B}_0^{-1}$$

Estimated structural IRF with 95% CI (Fig 2 in Galí)

Figure 2 : US



- Conclusion: technology shock raises productivity and lowers hours worked

$$\hat{\Omega} = \begin{bmatrix} 0.43 & -0.06 \\ -0.06 & 0.42 \end{bmatrix}$$

small negative correlation between VAR resids

$$\lim_{s \rightarrow \infty} \begin{bmatrix} \frac{\partial x_{t+s}}{\partial \varepsilon_{1t}} & \frac{\partial x_{t+s}}{\partial \varepsilon_{2t}} \\ \frac{\partial n_{t+s}}{\partial \varepsilon_{1t}} & \frac{\partial n_{t+s}}{\partial \varepsilon_{2t}} \end{bmatrix} = [\hat{\Phi}(1)]^{-1} = \begin{bmatrix} 0.89 & -0.50 \\ 0.87 & 1.73 \end{bmatrix}$$

For u_{2t} to have zero long-run effect on x_t , it must be interpreted as something that moves ε_{1t} and ε_{2t} in the same direction.

For u_{1t} to be uncorrelated with u_{2t} , it must be interpreted as something that moves ε_{1t} and ε_{2t} in opposite directions.

So if technology shock raises productivity, it must increase ε_{1t} and decrease ε_{2t} .

$$\boldsymbol{\varepsilon}_t = \mathbf{B}_0^{-1} \mathbf{u}_t$$

$$\hat{\mathbf{B}}_0^{-1} = \begin{bmatrix} 0.59 & 0.30 \\ -0.37 & 0.53 \end{bmatrix}$$

$$\begin{bmatrix} \frac{\partial \varepsilon_{1t}}{\partial u_{1t}} & \frac{\partial \varepsilon_{1t}}{\partial u_{2t}} \\ \frac{\partial \varepsilon_{2t}}{\partial u_{1t}} & \frac{\partial \varepsilon_{2t}}{\partial u_{2t}} \end{bmatrix} = \begin{bmatrix} 0.59 & 0.30 \\ -0.37 & 0.53 \end{bmatrix}$$

$$\begin{aligned}
\lim_{s \rightarrow \infty} \begin{bmatrix} \frac{\partial x_{t+s}}{\partial u_{1t}} & \frac{\partial x_{t+s}}{\partial u_{2t}} \\ \frac{\partial n_{t+s}}{\partial u_{1t}} & \frac{\partial n_{t+s}}{\partial u_{2t}} \end{bmatrix} &= [\hat{\Phi}(1)]^{-1} \mathbf{B}_0^{-1} \\
&= \begin{bmatrix} 0.89 & -0.50 \\ 0.87 & 1.73 \end{bmatrix} \begin{bmatrix} 0.59 & 0.30 \\ -0.37 & 0.53 \end{bmatrix} \\
&= \begin{bmatrix} 0.71 & 0 \\ -0.14 & 1.18 \end{bmatrix}
\end{aligned}$$

$$\begin{bmatrix} \frac{\partial \varepsilon_{1t}}{\partial u_{1t}} & \frac{\partial \varepsilon_{1t}}{\partial u_{2t}} \\ \frac{\partial \varepsilon_{2t}}{\partial u_{1t}} & \frac{\partial \varepsilon_{2t}}{\partial u_{2t}} \end{bmatrix} = \begin{bmatrix} 0.59 & 0.30 \\ -0.37 & 0.53 \end{bmatrix}$$

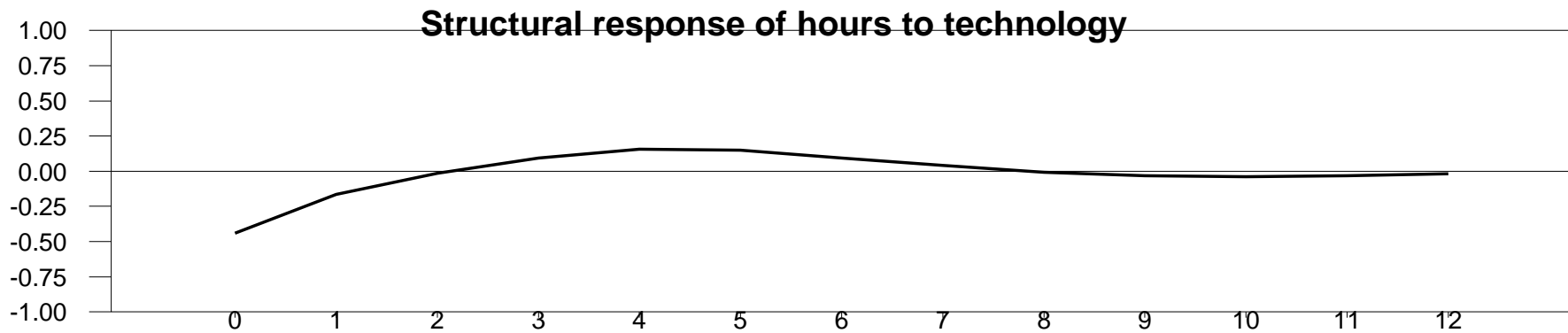
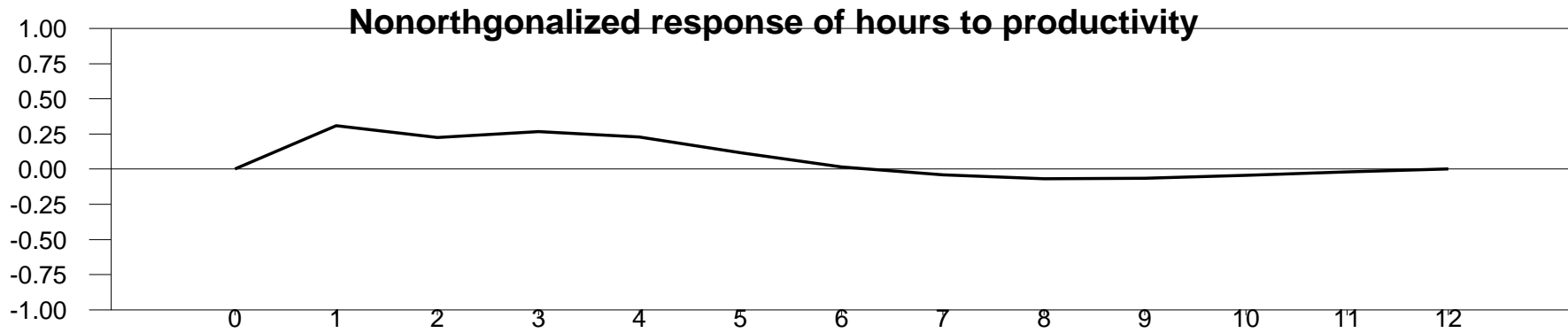
A one-unit shock to u_{2t} is interpreted as causing a 0.30 increase in ε_{1t} and 0.53 increase in ε_{2t}

$$\lim_{s \rightarrow \infty} \frac{\partial x_{t+s}}{\partial u_{2t}} = (0.89)(0.30) - (0.50)(0.53) = 0$$

$$\begin{bmatrix} \frac{\partial \varepsilon_{1t}}{\partial u_{1t}} & \frac{\partial \varepsilon_{1t}}{\partial u_{2t}} \\ \frac{\partial \varepsilon_{2t}}{\partial u_{1t}} & \frac{\partial \varepsilon_{2t}}{\partial u_{2t}} \end{bmatrix} = \begin{bmatrix} 0.59 & 0.30 \\ -0.37 & 0.53 \end{bmatrix}$$

A one-unit shock to u_{1t} is interpreted as causing a 0.59 increase in ε_{1t} and 0.37 decrease in ε_{2t}

$$\frac{\partial n_{t+s}}{\partial u_{1t}} = 0.59 \frac{\partial n_{t+s}}{\partial \varepsilon_{1t}} - 0.37 \frac{\partial n_{t+s}}{\partial \varepsilon_{2t}}$$



Additional comments:

$$(1) \hat{\mathbf{Q}} = \left(\mathbf{I}_2 - \hat{\mathbf{\Phi}}_1 - \hat{\mathbf{\Phi}}_2 - \cdots - \hat{\mathbf{\Phi}}_p \right)^{-1}$$

is estimated poorly, sensitive to p

Note: algorithm may be a little more robust if instead use

$$\hat{\mathbf{Q}} = \hat{\Psi}_0 + \hat{\Psi}_1 + \cdots + \hat{\Psi}_m \text{ for some } m$$

Chari, Kehoe and McGrattan (2008) give example of DSGE for which when this kind of procedure is applied does not uncover technology shock.

- (2) technology shock could be temporary
(e.g., delay in adoption of discovered technology)
- (3) demand shock could be permanent
(e.g., lost human capital)

C. Identification using heteroskedasticity (Wright, 2012)

t = daily Nov 3, 2008 to Sep 30, 2011

y_{1t} = 2-year Treasury yield

y_{2t} = 10-year Treasury yield

y_{3t} = 5-year TIPS break-even

(nominal yield minus TIPS yield)

y_{4t} = 5-10-year TIPS forward break-even

($2 \times$ 10-year TIPS break-even minus

5-year TIPS break-even)

y_{5t} = BAA yield

y_{6t} = AAA yield

$$\mathbf{y}_t = \mathbf{c} + \Phi_1 \mathbf{y}_{t-1} + \dots + \Phi_p \mathbf{y}_{t-p} + \boldsymbol{\varepsilon}_t$$

$$\boldsymbol{\varepsilon}_t = \mathbf{B}_0^{-1} \mathbf{u}_t$$

u_{1t} = monetary policy shock

want to estimate $\mathbf{b}^{(1)}$ (first
column of \mathbf{B}_0^{-1})

Suppose we believed that:

(1) monetary policy shocks have higher variance on particular days

$$E(u_{1t}^2) = \begin{cases} d_{11}^{(0)} + \lambda & \text{if } t \in S \\ d_{11}^{(0)} & \text{if } t \notin S \end{cases}$$

Set S is known (Wright uses FOMC dates and dates of monetary policy announcements)

Date	Event	Time
11/25/2008	<i>Fed announces purchases of MBS and agency bonds</i>	08:15
12/1/2008	<i>Bernanke states Treasuries may be purchased</i>	13:45
12/16/2008	<i>FOMC Meeting</i>	14:15
1/28/2009	<i>FOMC Meeting</i>	14:15
3/18/2009	<i>FOMC Meeting</i>	14:15
4/29/2009	FOMC Meeting	14:15
6/24/2009	FOMC Meeting	14:15
8/12/2009	FOMC Meeting	14:15
9/23/2009	FOMC Meeting	14:15
11/4/2009	FOMC Meeting	14:15
12/16/2009	FOMC Meeting	14:15
1/27/2010	FOMC Meeting	14:15
3/16/2010	FOMC Meeting	14:15
4/28/2010	FOMC Meeting	14:15
6/23/2010	FOMC Meeting	14:15
8/10/2010	<i>FOMC Meeting</i>	14:15
8/27/2010	<i>Bernanke speech at Jackson Hole</i>	10:00
9/21/2010	<i>FOMC Meeting</i>	14:15
10/15/2010	<i>Bernanke speech at Boston Fed</i>	08:15
11/3/2010	<i>FOMC Meeting</i>	14:15
12/14/2010	FOMC Meeting	14:15
1/26/2011	FOMC Meeting	14:15
3/15/2011	FOMC Meeting	14:15
4/27/2011	FOMC Meeting	12:30
6/2/2011	FOMC Meeting	12:30
8/9/2011	<i>FOMC Meeting</i>	14:15
8/26/2011	<i>Bernanke Speech at Jackson Hole</i>	10:00
9/21/2011	<i>FOMC Meeting</i>	14:15

(2) A monetary policy shock of given size would have the same effects on these dates as others

(3) Variance and effects of other shocks same on these dates as others

Then

$$E(\mathbf{u}_t \mathbf{u}_t') = \begin{cases} \mathbf{D} + \lambda \mathbf{e}_1 \mathbf{e}_1' & \text{if } t \in S \\ \mathbf{D} & \text{if } t \notin S \end{cases}$$

$\mathbf{e}_1 = \text{col 1 of } \mathbf{I}_n$

$$\boldsymbol{\varepsilon}_t = \mathbf{B}_0^{-1} \mathbf{u}_t = \sum_{i=1}^n \mathbf{b}^{(i)} u_{it}$$

$$E(\boldsymbol{\varepsilon}_t \boldsymbol{\varepsilon}_t') = \begin{cases} \mathbf{B}_0^{-1} \mathbf{D} (\mathbf{B}_0^{-1})' + \lambda \mathbf{b}^{(1)} (\mathbf{b}^{(1)})' & \text{if } t \in S \\ \mathbf{B}_0^{-1} \mathbf{D} (\mathbf{B}_0^{-1})' & \text{if } t \notin S \end{cases}$$

$$\hat{\mathbf{\Omega}}_1 = T_1^{-1} \sum_{t=1}^T \hat{\mathbf{\epsilon}}_t \hat{\mathbf{\epsilon}}_t' \delta(t \in S)$$

$$T_1 = \sum_{t=1}^T \delta(t \in S)$$

$$\hat{\mathbf{\Omega}}_0 = T_0^{-1} \sum_{t=1}^T \hat{\mathbf{\epsilon}}_t \hat{\mathbf{\epsilon}}_t' \delta(t \notin S)$$

$$T_0 = \sum_{t=1}^T \delta(t \notin S)$$

$$\hat{\mathbf{\Omega}}_1 - \hat{\mathbf{\Omega}}_0 \xrightarrow{p} \lambda \mathbf{b}^{(1)} (\mathbf{b}^{(1)})'$$

so we can estimate $\mathbf{b}^{(1)}$ up to an

unknown scale, e.g.: normalize $\lambda = 1$

$$\sqrt{T_1} [\text{vech}(\hat{\Omega}_1) - \text{vech}(\Omega_1)]$$

$$\xrightarrow{L} N(\mathbf{0}, \mathbf{V}_1)$$

element of \mathbf{V}_1 corresponding to
covariance between $\hat{\sigma}_{ij}$ and $\hat{\sigma}_{\ell m}$
given by $(\sigma_{i\ell}\sigma_{jm} + \sigma_{im}\sigma_{j\ell})$
(Hamilton, TSA, p. 301).

(1) Test null hypothesis that $\mathbf{\Omega}_0 = \mathbf{\Omega}_1$

$$\hat{\mathbf{q}}' [\hat{\mathbf{V}}_1/T_1 + \hat{\mathbf{V}}_0/T_0]^{-1} \hat{\mathbf{q}} \xrightarrow{L} \chi^2(n(n+1)/2)$$

$$\hat{\mathbf{q}} = \text{vech}(\hat{\mathbf{\Omega}}_1) - \text{vech}(\hat{\mathbf{\Omega}}_0)$$

or bootstrap critical value

This statistic is 67.9.

asymptotic: $P[\chi^2(21) > 46.8] = 0.001$

bootstrap p -value = 0.005

\Rightarrow reject $H_0 : \Omega_0 = \Omega_1$

Variance on announcement days different from others so this assumption of framework is correct.

(2) Estimate $\mathbf{b}^{(1)}$ by minimum chi square:

$$\hat{\mathbf{b}}^{(1)} = \arg \min_{\mathbf{b}^{(1)}} \tilde{\mathbf{q}}' [\hat{\mathbf{V}}_1/T_1 + \hat{\mathbf{V}}_0/T_0]^{-1} \tilde{\mathbf{q}}$$

$$\tilde{\mathbf{q}} = \hat{\mathbf{q}} - \text{vech}[\mathbf{b}^{(1)} (\mathbf{b}^{(1)})']$$

$$\widehat{\frac{\partial \mathbf{y}_{t+s}}{\partial u_{1t}}} = \hat{\Psi}_s \hat{\mathbf{b}}^{(1)}$$

(3) Test null hypothesis restriction valid:
value of objective function asymptotically
 $\chi^2(n(n-1)/2)$ or bootstrap critical value.

This test does not reject

\Rightarrow assumption that $\mathbf{\Omega}_1 = \mathbf{\Omega}_0 + \mathbf{b}^{(1)}\mathbf{b}^{(1)'}$

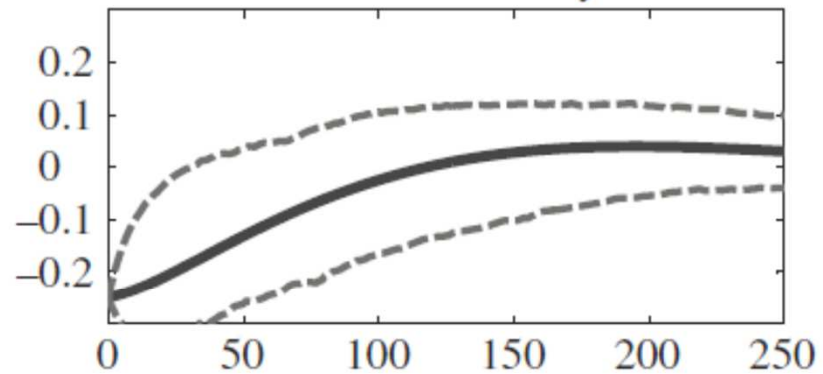
is consistent with observed data.

Normalization: second element of

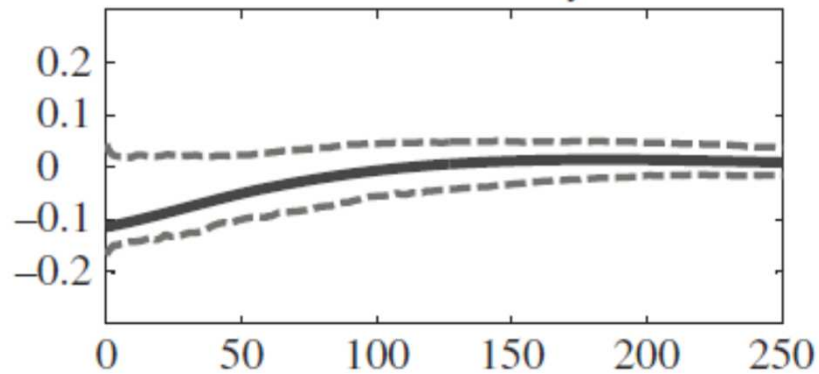
$$\mathbf{b}^{(1)} = -0.25$$

Monetary policy shock lowers 10-year yield by 25 bp.

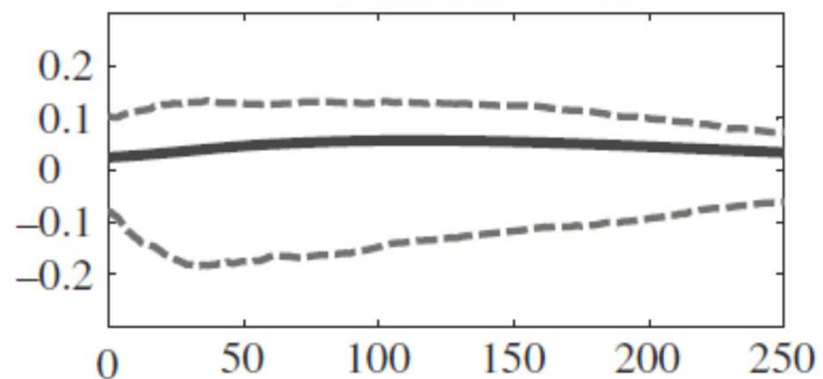
10 Year Treasury



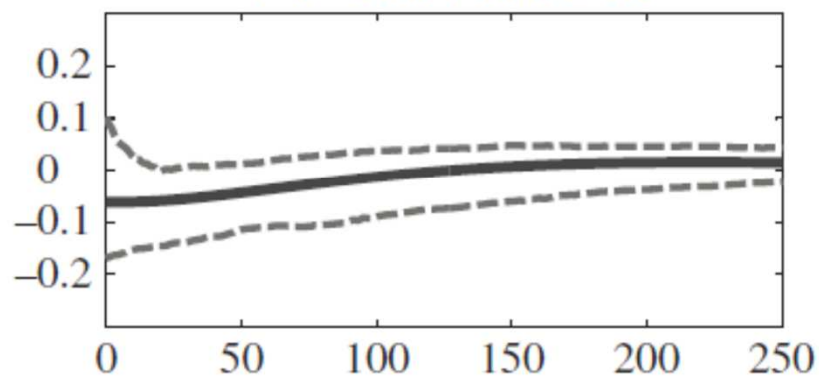
2 Year Treasury



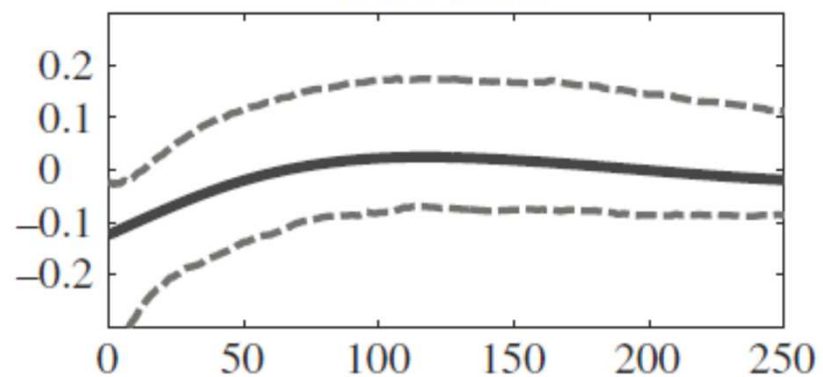
5 Year Break-even



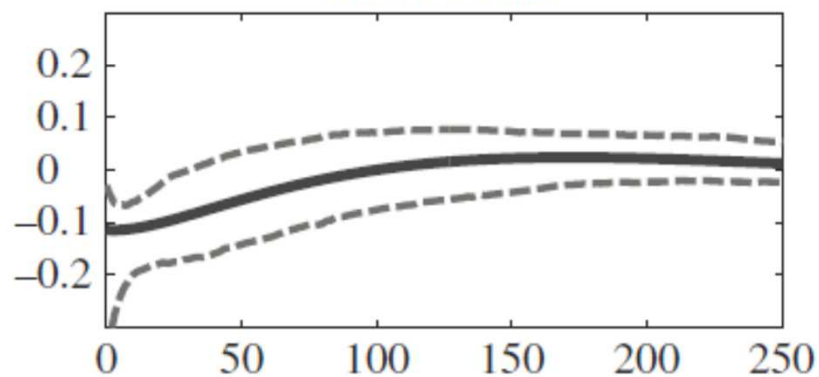
5-10 Year Break-even



BAA Yields



AAA Yields



- No evidence that unconventional monetary policy works through changes in expected inflation or risk premia.