## Orthogonalized VARs

A. Recursively orthogonalized VAR
B. Variance decomposition
C. Historical decomposition
D. Structural interpretation
E. Generalized IRFs

## A. Recursively orthogonalized VARs

Nonorthogonal IRF:

$$
\underset{(n \times n)}{\mathbf{\Psi}_{s}}=\frac{\partial E\left(\mathbf{y}_{t+s} \mid \mathbf{y}_{t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial \mathbf{y}_{t}^{\prime}}
$$

Column $1=\frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{1 t}, y_{2 t}, \ldots, y_{n t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{1 t}}$
e.g., already have data on $y_{2 t}, \ldots, y_{n t}$ and ask how a 1-unit change in $y_{1 t}$ affects forecast.

Could instead ask $\frac{\partial E\left(y_{t+t} y_{l} y_{1}, y_{t-1}, \ldots, y_{t-p}\right)}{\partial y_{1 t}}$
e.g., don't have any data from period $t$ except for $y_{1 t}$ and ask how 1 -unit change in $y_{1 t}$ affects forecast.
Knowing $y_{1 t}$ gives us information about $y_{2 t}, \ldots, y_{n t}$ if VAR forecast errors are correlated.

How calculate $\underset{(n \times 1)}{\mathbf{h}_{s 1}}=\frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{1 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{1 t}}$ ?
Method 1: local projection
Estimate by $n$ OLS equations

$$
\mathbf{y}_{t+s}=\mathbf{c}_{s}+\mathbf{h}_{s 1} y_{1 t}+\mathbf{H}_{s 2} \mathbf{y}_{t-1}+\cdots+\mathbf{H}_{s p} \mathbf{y}_{t-p+1}+\mathbf{u}_{t+s}
$$

Method 2: calculate answer implied by VAR

$$
\begin{aligned}
& \underset{(n \times 1)}{\mathbf{y}_{t}}=\left(y_{1 t}, y_{2 t}, \ldots, y_{n t}\right)^{\prime} \\
& \underset{(k \times 1)}{\mathbf{x}_{t}}=\left(1, \mathbf{y}_{t-1}^{\prime}, \mathbf{y}_{t-2}^{\prime}, \ldots, \mathbf{y}_{t-p}^{\prime}\right)^{\prime} \\
& k=n p+1 \\
& \mathbf{y}_{t}=\Gamma^{\prime} \mathbf{x}_{t}+\boldsymbol{\varepsilon}_{t} \\
& E\left(\boldsymbol{\varepsilon}_{t} \boldsymbol{\varepsilon}_{t}^{\prime}\right)=\boldsymbol{\Omega}
\end{aligned}
$$

Given parameters, observation of $y_{1 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}$ allows us to observe
$\varepsilon_{1 t}=y_{1 t}-\gamma_{1}^{\prime} \mathbf{x}_{t}$

Can calculate optimal forecast of $\varepsilon_{i t}$ given $\varepsilon_{1 t}$ as $E\left(\varepsilon_{i t} \mid \varepsilon_{1 t}\right)=\frac{\sigma_{i l}}{\sigma_{11}} \varepsilon_{1 t}$
$E\left(\boldsymbol{\varepsilon}_{t} \mid \varepsilon_{1 t}\right)=\left[\begin{array}{c}1 \\ \sigma_{21} / \sigma_{11} \\ \vdots \\ \sigma_{n 1} / \sigma_{11}\end{array}\right] \varepsilon_{1 t}=\mathbf{a}_{1} \varepsilon_{1 t}$

$$
\begin{aligned}
& \frac{\partial E\left(\mathbf{y}_{t} \mid y_{1 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{1 t}}=\mathbf{a}_{1} \\
& \frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{1 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{1 t}}=\frac{\partial E\left(\mathbf{y}_{t+s} \mid \mathbf{y}_{t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial \mathbf{y}_{t}^{\prime}} \frac{\partial E\left(\mathbf{y}_{t} \mid y_{1 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{1 t}} \\
&=\Psi_{s} \mathbf{a}_{1}
\end{aligned}
$$

$$
\hat{\Gamma}(n \times k)_{\hat{\Gamma}^{\prime}}=\left(\sum_{t=1}^{T} \mathbf{y}_{t} \mathbf{x}_{t}^{\prime}\right)\left(\sum_{t=1}^{T} \mathbf{x}_{\mathbf{x}} \mathbf{x}_{t}^{\prime-1} \Rightarrow \hat{\Psi}_{s}\right.
$$

$$
\hat{\boldsymbol{\varepsilon}}_{t}=\mathbf{y}_{t}-\hat{\Gamma}^{\prime} \mathbf{x}_{t}
$$

$$
(n \times 1)
$$

$$
\underset{(n \times n)}{\hat{\Omega}}=T^{-1} \sum_{t=1}^{T} \hat{\varepsilon}_{t} \hat{\varepsilon}_{t}^{\prime} \Rightarrow \hat{\mathbf{a}}_{1}=\left[\begin{array}{c}
1 \\
\hat{\sigma}_{21} / \hat{\sigma}_{11} \\
\vdots \\
\hat{\sigma}_{n 1} / \hat{\sigma}_{11}
\end{array}\right]
$$

Could also do this using Cholesky factor:

$$
\begin{aligned}
& \hat{\mathbf{\Omega}}=\hat{\mathbf{P}} \hat{\mathbf{P}}^{\prime} \quad(\hat{\mathbf{P}} \text { lower triangular) } \\
& \hat{\mathbf{a}}_{1}=\hat{\mathbf{p}}_{1} \hat{p}_{11} \\
& \hat{\mathbf{p}}_{1}=\text { column } 1 \text { of } \hat{\mathbf{P}} \\
& \hat{p}_{11}=\text { row } 1 \text { col } 1 \text { element of } \hat{\mathbf{P}}
\end{aligned}
$$

$\Psi_{s} \mathbf{a}_{1}=\frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{1 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{1 t}}$
is effect of one-unit increase in $y_{1 t}$ or $\varepsilon_{1 t}$ on forecast $\Psi_{s} \mathbf{p}_{1}$ is effect of one-standard-deviation increase in $\varepsilon_{1 t}$ on forecast.
$\Psi_{s} \mathbf{a}_{1}$ shows IRF to shock in observed units $\Psi_{s} \mathbf{p}_{1}$ shows IRF to shock of typical size Plots will look identical just with different units on vertical axis $\Psi_{s} \mathbf{p}_{1}=\Psi_{s} \mathbf{a}_{1} p_{11}$

## Recursively orthgonalized VAR estimated 1954-2007



## Recursively orthgonalized local projection estimated 1954-2007





Suppose next that we've observed $y_{1 t}$ and $y_{2 t}$ but not $y_{3 t}, y_{4 t}, \ldots, y_{n t}$.
What is effect on forecast of changing $y_{2 t}$ ?
$\frac{\partial E\left(\mathbf{y}_{t} \mid y_{1 t}, y_{2 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{2 t}}=\mathbf{a}_{2}=\mathbf{p}_{2} / p_{22}$
$\mathbf{p}_{2}=$ column 2 of $\mathbf{P}$
$p_{22}=$ row 2 col 2 element of $\mathbf{P}$
first element of $\mathbf{a}_{2}$ is zero

$$
\begin{aligned}
& \frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{1 t}, y_{2 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{2 t}}=\frac{\partial E\left(\mathbf{y}_{t+s} \mid \mathbf{y}_{t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial \mathbf{y}_{t}^{\prime}} \frac{\partial E\left(\mathbf{y}_{t} \mid y_{1 t}, y_{2 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{2 t}} \\
& \quad=\Psi_{s} \mathbf{a}_{2}
\end{aligned}
$$


( $n \times n$ ) matrix of recursively orthogonalized shocks:

$$
\Psi_{s} \mathbf{P} \text { or } \Psi_{s} \mathbf{A}
$$

$$
\mathbf{A}=\left[\begin{array}{cccc}
1 & 0 & \cdots & 0 \\
a_{21} & 1 & \cdots & 0 \\
\vdots & \vdots & \cdots & \vdots \\
a_{n 1} & a_{n 2} & \cdots & 1
\end{array}\right]=\mathbf{P}[\operatorname{diag}(\mathbf{P})]^{-1}
$$

Note last col of $\Psi_{s} \mathbf{A}$ is identical to last col of $\Psi_{s}$

We have broken down the news arriving in period $t$ into $n$ separate uncorrelated components
$\varepsilon_{1 t}=y_{1 t}-\gamma_{1}^{\prime} \mathbf{x}_{t}=$ news about $y_{1 t}$
$u_{2 t}=\varepsilon_{2 t}-a_{21} \varepsilon_{1 t}=$ news about $y_{2 t}$ not already revealed by $y_{1 t}$

$$
\begin{aligned}
u_{n t} & =\varepsilon_{n t}-E\left(\varepsilon_{n t} \mid \varepsilon_{1 t}, \ldots, \varepsilon_{n-1, t}\right) \\
& =\text { news about } y_{n t} \text { not already revealed by }
\end{aligned}
$$

$y_{1 t}, \ldots, y_{n-1, t}$

Simple way to summarize these components:

$$
\begin{aligned}
& \underset{(n \times 1)}{\mathbf{v}_{t}}=\mathbf{P}^{-1} \boldsymbol{\varepsilon}_{t}=\left[\begin{array}{cccc}
p^{11} & 0 & \cdots & 0 \\
p^{21} & p^{22} & \cdots & 0 \\
\vdots & \vdots & \cdots & \vdots \\
p^{n 1} & p^{n 2} & \cdots & p^{n n}
\end{array}\right]\left[\begin{array}{c}
\varepsilon_{1 t} \\
\varepsilon_{2 t} \\
\vdots \\
\varepsilon_{n t}
\end{array}\right] \\
& E\left(\mathbf{v}_{t} \mathbf{v}_{t}^{\prime}\right)=\mathbf{P}^{-1} E\left(\boldsymbol{\varepsilon}_{t} \boldsymbol{\varepsilon}_{t}^{\prime}\right) \mathbf{P}^{\prime-1}=\mathbf{P}^{-1} \Omega \mathbf{P}^{\prime-1} \\
& =\mathbf{P}^{-1} \mathbf{P} \mathbf{P}^{\prime} \mathbf{P}^{\prime-1}=\mathbf{I}_{n}
\end{aligned}
$$

$\mathbf{v}_{t}$ is a linear combination of $\varepsilon_{t}$ whose elements are uncorrelated with each other

$$
\begin{aligned}
v_{1 t}= & p^{11} \varepsilon_{1 t} \\
& =\text { rescaled error forecasting } y_{1 t} \\
v_{2 t}= & p^{21} \varepsilon_{1 t}+p^{22} \varepsilon_{2 t} \\
& =\text { rescaled error forecasting } \varepsilon_{2 t} \text { from } \varepsilon_{1 t} \\
v_{n t}= & p^{n 1} \varepsilon_{1 t}+p^{n 2} \varepsilon_{2 t}+\cdots+p^{n n} \varepsilon_{n t} \\
& =\text { rescaled error forecasting } \varepsilon_{n t} \text { from } \varepsilon_{1 t}, \ldots, \varepsilon_{n-1, t}
\end{aligned}
$$

## B. Variance decomposition

$$
\begin{aligned}
& \mathbf{y}_{t+s}=\hat{\mathbf{y}}_{t+s \mid t}+\Psi_{0} \boldsymbol{\varepsilon}_{t+s}+\Psi_{1} \boldsymbol{\varepsilon}_{t+s-1}+\Psi_{2} \boldsymbol{\varepsilon}_{t+s-2}+\cdots+\Psi_{s-1} \boldsymbol{\varepsilon}_{t+1} \\
& E\left(\mathbf{y}_{t+s}-\hat{\mathbf{y}}_{t+s \mid t}\right)\left(\mathbf{y}_{t+s}-\hat{\mathbf{y}}_{t+s \mid t}\right)=\sum_{m=0}^{s-1} \boldsymbol{\Psi}_{m} \boldsymbol{\Omega} \Psi_{m}^{\prime} \\
& \boldsymbol{\varepsilon}_{t}=\mathbf{P} \mathbf{v}_{t}=\mathbf{p}_{1} v_{1 t}+\mathbf{p}_{2} v_{2 t}+\cdots+\mathbf{p}_{n} v_{n t}
\end{aligned}
$$

Contribution of $v_{i, t+1}, v_{i, t+2}, \ldots, v_{i, t+s}$ to forecast error:
$\Psi_{0} \mathbf{p}_{i} v_{i, t+s}+\Psi_{1} \mathbf{p}_{i} v_{i, t+s-1}+\cdots+\boldsymbol{\Psi}_{s-1} \mathbf{p}_{i} v_{i, t+1}$

$$
\begin{aligned}
& E\left(\mathbf{y}_{t+s}-\hat{\mathbf{y}}_{t+s|l|}\right)\left(\mathbf{y}_{t+s}-\hat{\mathbf{y}}_{t+s \mid t}\right) \\
& \quad=\sum_{m=0}^{s-1} \Psi_{m} \mathbf{p}_{1} \mathbf{p}_{1}^{\prime} \Psi_{m}^{\prime}+\cdots+\sum_{m=0}^{s-1} \Psi_{m} \mathbf{p}_{n} \mathbf{p}_{n}^{\prime} \Psi_{m}^{\prime}
\end{aligned}
$$

First term: amount by which could reduce MSE if we knew the values of $\varepsilon_{1, t+1}, \ldots, \varepsilon_{1, t+s}$
Second term: amount by which we could reduce
MSE if we knew the values of $u_{2, t+1}, \ldots, u_{2, t+s}$

Decomposition of Variance for Series GDPCH Step Std Error GDPCH INFLATION FEDFUNDS

| 1 | 3.08741913 | 100.000 | 0.000 | 0.000 |
| ---: | ---: | ---: | ---: | ---: |
| 2 | 3.15963628 | 99.994 | 0.001 | 0.005 |
| 3 | 3.36710884 | 90.098 | 0.218 | 9.684 |
| 4 | 3.41017407 | 87.881 | 1.147 | 10.972 |
| 5 | 3.43030191 | 87.226 | 1.928 | 10.846 |
| 6 | 3.45041262 | 86.487 | 2.627 | 10.886 |
| 7 | 3.45981410 | 86.245 | 2.871 | 10.883 |
| 8 | 3.46834423 | 86.053 | 3.093 | 10.854 |
| 9 | 3.47514728 | 85.850 | 3.331 | 10.819 |
| 10 | 3.47997093 | 85.727 | 3.483 | 10.789 |
| 11 | 3.48341375 | 85.612 | 3.612 | 10.776 |
| 12 | 3.48663478 | 85.487 | 3.744 | 10.769 |
| 13 | 3.48920625 | 85.388 | 3.856 | 10.756 |
| 14 | 3.49137054 | 85.302 | 3.954 | 10.745 |
| 15 | 3.49341827 | 85.216 | 4.047 | 10.737 |
| 16 | 3.49526789 | 85.139 | 4.132 | 10.729 |
| 17 | 3.49686734 | 85.072 | 4.206 | 10.721 |
| 18 | 3.49830813 | 85.011 | 4.273 | 10.716 |
| 19 | 3.49960777 | 84.956 | 4.333 | 10.711 |
| 20 | 3.50074296 | 84.907 | 4.385 | 10.708 |


| 1 | 1.12790671 | 0.711 | 99.289 | 0.000 |
| ---: | :--- | :--- | :--- | :--- |
| 2 | 1.30103516 | 1.009 | 97.570 | 1.421 |
| 3 | 1.43945588 | 1.909 | 96.580 | 1.511 |
| 4 | 1.58901603 | 1.634 | 97.023 | 1.343 |
| 5 | 1.73329210 | 4.906 | 93.947 | 1.147 |
| 6 | 1.83842144 | 6.087 | 92.878 | 1.034 |
| 7 | 1.93358191 | 6.981 | 91.717 | 1.302 |
| 8 | 2.01323981 | 7.629 | 90.802 | 1.569 |
| 9 | 2.07617743 | 7.983 | 90.376 | 1.641 |
| 10 | 2.12893320 | 8.188 | 89.988 | 1.824 |
| 11 | 2.17571089 | 8.286 | 89.663 | 2.051 |
| 12 | 2.21501613 | 8.322 | 89.481 | 2.197 |
| 13 | 2.24826561 | 8.329 | 89.336 | 2.335 |
| 14 | 2.27737512 | 8.303 | 89.209 | 2.488 |
| 15 | 2.30262070 | 8.268 | 89.109 | 2.623 |
| 16 | 2.32420833 | 8.229 | 89.031 | 2.740 |
| 17 | 2.34293779 | 8.186 | 88.958 | 2.856 |
| 18 | 2.35921378 | 8.141 | 88.891 | 2.968 |
| 19 | 2.37321184 | 8.097 | 88.834 | 3.069 |
| 20 | 2.38526891 | 8.056 | 88.782 | 3.163 |


| Decomposition of Variance for Series FEDFUNDS |  |  |  |  |
| ---: | :---: | ---: | ---: | ---: |
| Step | Std Error | GDPCH | INFLATION | FEDFUNDS |
| 1 | 0.78294284 | 3.309 | 2.939 | 93.752 |
| 2 | 1.24088336 | 11.710 | 8.195 | 80.095 |
| 3 | 1.51928430 | 20.625 | 13.752 | 65.623 |
| 4 | 1.76776027 | 25.568 | 18.317 | 56.116 |
| 5 | 1.97194315 | 28.715 | 20.684 | 50.601 |
| 6 | 2.13303190 | 31.520 | 22.672 | 45.807 |
| 7 | 2.27057752 | 33.433 | 24.707 | 41.860 |
| 8 | 2.39054533 | 34.802 | 26.385 | 38.813 |
| 9 | 2.49021257 | 35.694 | 27.902 | 36.404 |
| 10 | 2.57500835 | 36.215 | 29.393 | 34.392 |
| 11 | 2.64925525 | 36.494 | 30.775 | 32.732 |
| 12 | 2.71414804 | 36.608 | 32.042 | 31.350 |
| 13 | 2.77123546 | 36.605 | 33.230 | 30.165 |
| 14 | 2.82208295 | 36.523 | 34.335 | 29.142 |
| 15 | 2.86734779 | 36.391 | 35.348 | 28.261 |
| 16 | 2.90764045 | 36.226 | 36.275 | 27.499 |
| 17 | 2.94364156 | 36.040 | 37.125 | 26.836 |
| 18 | 2.97580221 | 35.843 | 37.897 | 26.260 |
| 19 | 3.00448248 | 35.642 | 38.597 | 25.761 |
| 20 | 3.03006847 | 35.441 | 39.230 | 25.329 |

## C. Historical decomposition

$$
\begin{aligned}
\mathbf{y}_{t+s} & =\hat{\mathbf{y}}_{t+s \mid t}+\sum_{m=0}^{s-1} \boldsymbol{\Psi}_{m} \boldsymbol{\varepsilon}_{t+s-m} \\
& =\hat{\mathbf{y}}_{t+s \mid t}+\sum_{m=0}^{s-1} \boldsymbol{\Psi}_{m}\left[\mathbf{p}_{1} v_{1, t+s-m}+\cdots+\mathbf{p}_{n} v_{n, t+s-m}\right]
\end{aligned}
$$

Can decompose the observed value for any variable at any date into component that could have been predicted as of some earlier date plus innovations in individual $v_{i, t+m}$ since then.

Historical Decomposition of GDP


## D. Structural interpretation

Suppose we hypothesized the following structural model for the behavior of the Fed:
$i_{t}=\lambda_{3}+\psi_{y} y_{t}+\psi_{\pi} \pi_{t}+\mathbf{b}_{31}^{\prime} \mathbf{y}_{t-1}+\cdots+\mathbf{b}_{3 p}^{\prime} \mathbf{y}_{t-p}+u_{3 t}$
$i_{t}=$ fed funds rate
$y_{t}=$ GDP growth rate
$\pi_{t}=$ inflation rate
$\psi_{y}, \psi_{\pi}=$ coefficients in Taylor Rule
$\mathbf{b}_{3 m}$ allow for inertia in monetary policy
$u_{3 t}=$ serially uncorrelated shock to monetary policy = deviation from Fed's usual rule, uncorrelated with $\mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}$ by definition
Would like to know $\partial \mathbf{y}_{t+s} / \partial u_{3 t}$

Suppose I also thought there was a Phillips Curve of the form
$\pi_{t}=\lambda_{2}+\alpha y_{t}+\mathbf{b}_{21}^{\prime} \mathbf{y}_{t-1}+\cdots+\mathbf{b}_{2 p}^{\prime} \mathbf{y}_{t-p}+u_{2 t}$
$\alpha=$ slope of Phillips Curve
$\mathbf{b}_{2 m}$ allow for inertia in PC
$u_{2 t}=$ unpredictable shock to PC
$u_{2 t}$ uncorrelated with $\mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}$ by definition
$u_{2 t}$ also assumed to be uncorrelated with $u_{3 t}$ (assumption that monetary policy shocks take more than one period to affect inflation)

Model equilibrium output as
$y_{t}=\lambda_{1}+\mathbf{b}_{11}^{\prime} \mathbf{y}_{t-1}+\cdots+\mathbf{b}_{1 p}^{\prime} \mathbf{y}_{t-p}+u_{1 t}$
$u_{1 t}=$ error forecasting GDP one period ahead
$u_{1 t}$ uncorrelated with $\mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}$ by definition
$u_{1 t}$ also assumed to be uncorrelated with $u_{2 t}, u_{3 t}$ (assumption that PC and monetary shocks take more than one period to affect output)
$i_{t}=\lambda_{3}+\psi_{y} y_{t}+\psi_{\pi} \pi_{t}+\mathbf{b}_{31}^{\prime} \mathbf{y}_{t-1}+\cdots+\mathbf{b}_{3 p}^{\prime} \mathbf{y}_{t-p}+u_{3 t}$
Above assumptions mean $u_{3 t}$ uncorrelated
with $y_{t}$ and $\pi_{t}$.
$\Rightarrow$ could estimate by OLS
$\hat{\psi}_{y}$ and $\hat{\psi}_{\pi}$ are same as step 0 Jordá projection $\hat{\psi}_{y}$ and $\hat{\psi}_{\pi}$ are same as $\hat{a}_{31}$ and $\hat{a}_{32}$
$\pi_{t}=\lambda_{2}+\alpha y_{t}+\mathbf{b}_{21}^{\prime} \mathbf{y}_{t-1}+\cdots+\mathbf{b}_{2 p}^{\prime} \mathbf{y}_{t-p}+u_{2 t}$
Above assumptions mean $u_{2 t}$ uncorrelated with $y_{t}$.
$\Rightarrow$ could estimate by OLS
$\hat{\alpha}$ is same as step 0 Jordá projection
$\hat{\alpha}$ is same as $\hat{a}_{21}$

Conclusion: under above assumptions with
$\mathbf{A}=\mathbf{P}[\operatorname{diag}(\mathbf{P})]^{-1}$
$\mathbf{u}_{t}=\mathbf{A}^{-1} \boldsymbol{\varepsilon}_{t}$
$u_{1 t}=\varepsilon_{1 t}$
The error I make forecasting $y_{1 t}$ given
$\mathbf{y}_{t-1}, \mathbf{y}_{t-2}, \ldots, \mathbf{y}_{t-p}$ is the shock to equilibrium output.

The error I make forecasting $y_{2 t}$ given $y_{1 t}, \mathbf{y}_{t-1}, \mathbf{y}_{t-2}, \ldots, \mathbf{y}_{t-p}$ is the shock to PC.

The error I make forecasting $y_{3 t}$ given
$y_{1 t}, y_{2 t}, \mathbf{y}_{t-1}, \mathbf{y}_{t-2}, \ldots, \mathbf{y}_{t-p}$ is the shock to monetary policy.

$$
\frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{1 t}, y_{2 t}, y_{3 t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{3 t}}=\frac{\partial \mathbf{y}_{t+s}}{\partial u_{3 t}}
$$

Recursively orthogonalized VAR gives the dynamic effects of monetary policy.

## Orthogonal Cholesky IRF with 95\%











- A monetary contraction (higher fed funds rate) is followed by slower GDP growth 2-3 quarters later
- But unanticipated monetary policy shocks account for only $10 \%$ of variance of output
- Most of variation in fed funds rate comes from predictable response of monetary policy to output and inflation
- A monetary contraction is followed by higher inflation (known as "price puzzle") ${ }^{38}$
- Assumption-free statement of price puzzle:
- if you tell me that fed funds rate is higher than

I would have predicted given current output, inflation, and lags, then I will revise my expectation of future inflation up.

- Natural interpretation:
- Fed raised funds rate because it anticipated future inflation.
- Our 3-variable equation is too simplistic a description of Fed
- Popular "fix" for price puzzle:
- Add other variables that better capture information about future inflation (such as commodity prices) to Fed policy equation

Christiano, Eichenbaum, Evans (1996)
$y_{1 t}=\log$ of real GDP
$y_{2 t}=\log$ of GDP deflator
$y_{3 t}=$ index of sensitive commodity prices
$y_{4 t}=$ fed funds rate
$y_{5 t}=$ nonborrowed reserves
$y_{6 t}=$ total reserves
$y_{7 t}=$ one of a set of macro variables

Structural model:
$\mathbf{B}_{0} \mathbf{y}_{t}=\mathbf{B} \mathbf{x}_{t}+\mathbf{u}_{t}$
$\mathbf{x}_{t}=\left(1, \mathbf{y}_{t-1}^{\prime}, \mathbf{y}_{t-2}^{\prime}, \ldots, \mathbf{y}_{t-p}^{\prime}\right)^{\prime}$
$\mathbf{u}_{t}=$ vector of structural shocks
$E\left(\mathbf{u}_{t} \mathbf{u}_{t}^{\prime}\right)=\mathbf{D}$ (diagonal)

$$
\mathbf{B}_{0}=\left[\begin{array}{lllllll}
1 & 0 & 0 & 0 & 0 & 0 & 0 \\
x & 1 & 0 & 0 & 0 & 0 & 0 \\
x & x & 1 & 0 & 0 & 0 & 0 \\
x & x & x & 1 & 0 & 0 & 0 \\
x & x & x & x & 1 & 0 & 0 \\
x & x & x & x & x & 1 & 0 \\
x & x & x & x & x & x & 1
\end{array}\right]
$$

Variable 4 is fed funds rate, equation 4 is monetary policy equation.

Note that

$$
\frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{1 t}, y_{2 t}, y_{3 t}, y_{4 t}, \mathbf{y}_{t-1}, \mathbf{y}_{t-2}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{4 t}}=\frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{2 t}, y_{1 t}, y_{3 t}, y_{4 t}, \mathbf{y}_{t-1}, \mathbf{y}_{t-2}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{4 t}}
$$

Will have the identical answer for effect of
variable 4 any way we order variables 1-3 and any way we order variables 5-7. Jordá estimate identical if reorder (keeping 4 in place).

If all we care about is effect of monetary policy, we only need to assume block-recursive
$\mathbf{B}_{0}=\left[\begin{array}{lllllll}x & x & x & 0 & 0 & 0 & 0 \\ x & x & x & 0 & 0 & 0 & 0 \\ x & x & x & 0 & 0 & 0 & 0 \\ x & x & x & 1 & 0 & 0 & 0 \\ x & x & x & x & x & x & x \\ x & x & x & x & x & x & x \\ x & x & x & x & x & x & x\end{array}\right]$


67\% confidence bands

## E. Generalized IRFs

- If we put fed funds fourth, estimated effect of monetary policy does not depend on how we order variables 1-3.
- But if we switch fed funds from 4 to 3 , results could change
- Pesaran and Shin (1998): "generalized impulse-response function"
- Put variable \#1 first to find effect of variable 1
- Put variable \#2 first to find effect of variable 2
- Put variable \#n first to find effect of variable n


## GIRF: for every $i$, calculate

$$
\frac{\partial E\left(\mathbf{y}_{t+s} \mid y_{i t}, \mathbf{y}_{t-1}, \ldots, \mathbf{y}_{t-p}\right)}{\partial y_{i t}}
$$

- Conclusion: any IRF or GIRF is giving answer to a forecasting question.
- Best practice: describe forecasting question explicitly and explain the reason that question is interesting.

