Clustered Standard Errors

- The Attraction of "Differences in Differences"
- 2. Grouped Errors Across Individuals
- Serially Correlated Errors

1. The Attraction of Differences in Differences Estimates

- Typically evaluate programs which differ across groups, such as U.S. States e.g., effect of changes in state minimum wage laws or state welfare programs on earnings or unemployment
- Treat selection (heterogeneity) bias by removing state effects (one "diff")
- Treat common economic fluctuations by removing year effects (the other "diff")
- Hence the appealing nickname "diffs in diffs"

2. The Grouped Error Problem:

- Binary covariates define groups within which errors are potentially correlated (e.g., cities, states, years, states after treatment, self-employed, etc..)
 - remember that errors contain unobserved variables
- $Y_{ist} = A_{st} + B_t + cX_{ist} + \beta I_{st} + \varepsilon_{ist},$
- s are groups (perhaps states)
- t is time
- I is an indicator for treatment, which occurs as the group x time level
- ε is an error term, which is not necessary iid.

2. Grouped Errors Across Individuals E.g., Minimum wages on NJ/Penn border

- Card and Krueger (1994) looked at the effects of minimum wages on employment in fast-food restaurants near the NJ – Penn border.
- Data collected before and after NJ raised its' minimum wage by 80 cents (in 1992).
- i restaurant, s state, t time
- S=2, T=2, N is large.
- They found small positive effects within a small confidence interval of zero.

2. Grouped Errors Across Individuals E.g., Mariel Boatlift

- Card (1990) looked at the effects of a surprise supply shock of immigrants to Miami due to a temporary lifting of emigration restrictions by Cuba in 1980.
- He estimates the effect of the boatlift on unemployment and wages of low skill workers in Miami using four other cities as comparisons (Atlanta, Houston, LA and Tampa-St. Petersburg) with CPS data.
- i individual, s city, t time
- S=5, T~=2, N is large.
- He finds no statistically significant effect on employment or wages of the labor supply shock.

2. Grouped Errors Across Individuals

- How big does the number of groups (S, or S*T) have to be?
- $Y_{ist} = a_{sb} + d_t + cZ_{ist} + \beta I_{st} + \varepsilon_{ist},$
- Donald and Lang (2004): In the (plausible) case where we have some within-group correlation, and under generous assumptions the t-statistics converge to a normal distribution at rate S*T no matter what N is.
- Intuition: Imagine that within s,t groups the errors are perfectly correlated. Then you might as well aggregate and run the regression with S*T observations.
- Intuition: 2 step estimator
- If group and time effects are included, with normally distributed group-time specific errors under generous assumptions, the tstatistics have a t distribution with S*T-S-T degrees of freedom, no matter what N is. (Table 3)
- Donald-Lang suggested estimator has this flavor. (Table 3)
- Alternative: collapse into s,t groups
- 3 issues: consistent s.e., efficient s.e. and distribution of t-stat in small samples

Distribution of t-ratio, 4 d.o.f, $\beta = 0$

TABLE 3 MONTE CARLO ESTIMATION Distribution of t-statistics (4 groups, 2500 observations per group)								
	99th percentile	95th percentile	90th percentile	% > 1.645	% > 1.96			
OLS (co	nventional standa	rd errors)						
No Z	13.01	9.93	8.40	74.5	69.8			
Z	13.01	9.93	8.40	74.5	69.9			
OLS (Ei	OLS (Eicker-White standard errors)							
No Z	15.56	6.74	4.58	39.6	33.0			
Z	14.58	6.74	4.58	39.6	33.0			
Feasible	Feasible GLS (random effects)							
No Z	7.34	4.00	2.82	23.8	18.6			
Z	9.74	6.07	4.30	32.7	27.4			
Two-Step								
No Z	9.72	4.28	2.92	24.1	18.9			
Z	9.75	4.28	2.92	24.1	18.9			

3. Correlations over time in panels

- $Y_{ist} = A_{st} + B_t + cX_{ist} + \beta I_{st} + \varepsilon_{ist},$
- S are groups (perhaps states)
- t is time
- I is an indicator for treatment, which occurs as the group x time level
- Correlations within group, period (i.e., s,t) cells only is very restrictive.
- In general we want to allow correlations over time as well (within s but not within t)

Lots of DD papers

T is large

The variables tend to be serially corr.

So are std.

errors consistent?

TABLE I SURVEY OF DD PAPERS^A

Number of DD papers	92	
Number with more than 2 periods of data	69	
Number which collapse data into before-after	4	
Number with potential serial correlation problem	65	
Number with some serial correlation correction	5	
GLS	4	
Arbitrary variance-covariance matrix	1	
Distribution of time span for papers with more than 2 periods	Average	16.5
	Percentile	Value
	1%	3
	5%	3
	10%	4
	25%	5.75
	50%	11
	75%	21.5
	90%	36
	95%	51
	99%	83
Most commonly used dependent variables	Number	
Employment	18	
Wages	13	
Health/medical expenditure	8	
Unemployment	6	
Fertility/teen motherhood	4	
Insurance	4	
Poverty	3	
Consumption/savings	3	
Informal techniques used to assess endogeneity	Number	
Graph dynamics of effect	15	
See if effect is persistent	2	
DDD	11	
Include time trend specific to treated states	7	
Look for effect prior to intervention	3	
Include lagged dependent variable	3	
Number with potential clustering problem	80	
Number which deal with it	36	

Data come from a survey of all articles in six journals between 1990 and 2000: the American Economic Review, the Industrial Labor Relations Review, the Journal of Labor Economics, the Journal of Political Economy, the Journal of Public Economics, and the Quarterly Journal of Economics. We define an article as "Difference-in-Difference" if it (1) examines the effect of a specific intervention and (2) uses units unaffected by the intervention as a control group.

Placebo Binary "Laws"

- Randomly choose a year between 79-99
 a randomly assign a law to 25 states tillend of 99
- Rej. rate is % for which t>1.96

TABLE II DD Rejection Rates for Placebo Laws

A. CPS DATA

			Rejection rate	
Data	$\hat{\rho}_1,\hat{\rho}_2,\hat{\rho}_3$	Modifications	No effect	2% effect
1) CPS micro, log			.675	.855
wage			(.027)	(.020)
CPS micro, log		Cluster at state-	.44	.74
wage		year level	(.029)	(.025)
CPS agg, log	.509, .440, .332		.435	.72
wage			(.029)	(.026)
4) CPS agg, log	.509, .440, .332	Sampling	.49	.663
wage		w/replacement	(.025)	(.024)
CPS agg, log	.509, .440, .332	Serially	.05	.988
wage		uncorrelated laws	(.011)	(.006)
CPS agg,	.470, .418, .367		.46	.88
employment			(.025)	(.016)
CPS agg, hours	.151, .114, .063		.265	.280
worked			(.022)	(.022)
8) CPS agg, changes	046, .032, .002		0	.978
in log wage	,,			(.007)

B. MONTE CARLO SIMULATIONS WITH SAMPLING FROM AR(1) DISTRIBUTION

			Rejecti	ection rate	
Data	ρ	Modifications	No effect	2% effect	
9) AR(1)	.8		.373	.725	
			(.028)	(.026)	
10) AR(1)	0		.053	.783	
			(.013)	(.024)	
11) AR(1)	.2		.123	.738	
			(.019)	(.025)	
12) AR(1)	.4		.19	.713	
			(.023)	(.026)	
13) AR(1)	.6		.333	.700	
			(.027)	(.026)	
14) AR(1)	4		.008	.7	
			(.005)	(.026)	

a. Unless mentioned otherwise under "Modifications," reported in the last two columns are the OLS rejection rates of the null hypothesis of no effect (at the 5 percent significance level) on the intervention variable for randomly generated placebo interventions as described in text. The data used in the last column were altered to simulate a true 2 percent effect of the intervention. The number of simulations for each cell is at least 200 and typically 400.

b. CPS data are data for women between 25 and 50 in the fourth interview month of the Merged Outgoing Rotation Group for the years 1979 to 1999. In rows 3 to 8 of Panel A, data are aggregated to state-year level cells after controlling for demographic variables (four education dummies and a quartic in age). For each simulation in rows 1 through 3, we use the observed CPS data. For each simulation in rows 4 through 8, the data generating process is the state-level empirical distribution of the CPS data that puts a probability of 1/50 on the different states' outcomes (see text for details). For each simulation in Panel B, the data generating process is an AR(1) model with normal disturbances chosen to match the CPS state female wage variances (see text for details). $\hat{\rho}_t$ refer to the estimated autocorrelation parameter of lag \hat{t} , p refers to the autocorrelation parameter in the AR(1) model.

c. All regressions include, in addition to the intervention variable, state and year fixed effects. The individual level regressions also include demographic controls.

Placebo Binary "Laws"

Type I error is worst when T is large

TABLE III Varying N and T

N 50 20	T A. CPS DATA 21 21	No effect .49 (.025)	2% effect
50	21		
20	21	(.025)	
20	21		(.024)
		.39	.54
		(.024)	(.025)
10	21	.443	.510
		(.025)	(.025)
6	21	.383	.433
		(.025)	(.025)
50	11	.20	.638
		(.020)	(.024)
50	7	.15	.635
		(.017)	(.024)
50	5	.078	.5
		(.013)	(.025)
50	3	.048	.363
		(.011)	(.024)
50	2	.055	.28
		(.011)	(.022)
IULATIONS	WITH SAMPLI	NG FROM AR(1) DIS	STRIBUTION
50	21	.35	.638
00			(.028)
20	21		.538
20			(.029)
10	21		.505
10			(.029)
6	21		.5
Ü	21		(.029)
50	11		.588
50	11		(.028)
50	5		.5525
50			(.029)
50	3		.435
50			(.029)
50	50		.855
50	00	(.029)	(.020)
	50 50 50 50	50 7 50 5 50 3 50 2 IULATIONS WITH SAMPLI 50 21 20 21 10 21 6 21 50 11 50 5 50 3	(.020) 50 7 .15 (.017) 50 5 .078 (.013) 50 3 .048 (.011) 50 2 .055 (.011) IULATIONS WITH SAMPLING FROM AR(1) DIS 50 21 .35 (.028) 20 21 .35 (.028) 20 21 .3975 (.028) 6 21 .393 (.028) 6 21 .393 (.028) 50 11 .335 (.027) 50 5 .175 (.022) 50 3 .09 (.017) 50 50 .4975

Solutions: AR(1) correction

- N=50, T=21
- AR(1) biased for small T
- Process looks more like AR(2)

TABLE IV
PARAMETRIC SOLUTIONS

			Rejection rate			
Data	Technique	Estimated $\hat{\rho}_1$	No effect	2% Effect		
	A. CPS I	DATA				
1) CPS aggregate	OLS		.49	.663		
			(.025)	(.024)		
2) CPS aggregate	Standard AR(1)	.381	.24	.66		
	correction		(.021)	(.024)		
3) CPS aggregate	AR(1) correction		.18	.363		
	imposing $\rho = .8$		(.019)	(.024)		
B. OTHER DATA GENERATING PROCESSES						
4) AR(1), $\rho = .8$	OLS		.373	.765		
, , , , , , , , , , , , , , , , , , , ,			(.028)	(.024)		
5) AR(1), $\rho = .8$	Standard AR(1)	.622	.205	.715		
	correction		(.023)	(.026)		
6) AR(1), $\rho = .8$	AR(1) correction		.06	.323		
	imposing $\rho = .8$		(.023)	(.027)		
7) AR(2), $\rho_1 = .55$	Standard AR(1)	.444	.305	.625		
$\rho_2 = .35$	correction		(.027)	(.028)		
8) $AR(1)$ + white	Standard AR(1)	.301	.385	.4		
noise, $\rho = .95$,	correction		(.028)	(.028)		
noise/signal = .13						

Solutions: Ignore TS Information

- correct size but loss of power
- Residual aggregation is a Frisch-Waugh exercise: first - regress on other variables, then - aggregate residuals before and after treatment

TABLE VI IGNORING TIME SERIES DATA

			Rejection rate	
Data	Technique		No effect	2% effec
	A. CPS DATA			
1) CPS agg	OLS	50	.49	.663
			(.025)	(.024)
2) CPS agg	Simple aggregation	50	.053	.163
- and			(.011)	(.018)
3) CPS agg	Residual aggregation	50	.058	.173
CPS agg, staggered laws	Residual aggregation	50	(.011)	(.019) .363
t) CFS agg, staggered laws	Kesiddai aggregation	30	.048	(.024)
5) CPS agg	OLS	20	.39	.54
5, C1 5 agg	0120	20	(.025)	(.025)
6) CPS agg	Simple aggregation	20	.050	.088
	1 33 3		(.011)	(.014)
7) CPS agg	Residual aggregation	20	.06	.183
			(.011)	(.019)
B) CPS agg, staggered laws	Residual aggregation	20	.048	.130
o and	O. C.		(.011)	(.017)
9) CPS agg	OLS	10	.443	.51
10) CPS agg	Simple aggregation	10	(.025)	(.025)
(b) CFS agg	Simple aggregation	10	.053 (.011)	.065 $(.012)$
11) CPS agg	Residual aggregation	10	.093	.178
11) CID agg	nesidual aggregation	10	(.014)	(.019)
12) CPS agg, staggered laws	Residual aggregation	10	.088	.128
			(.014)	(.017)
13) CPS agg	OLS	6	.383	.433
			(.024)	(.024)
14) CPS agg	Simple aggregation	6	.068	.07
(E) CDC	B :1 1		(.013)	(.013)
15) CPS agg	Residual aggregation	6	.11	.123
16) CPS agg, staggered laws	Residual aggregation	6	(.016) .09	(.016) .138
16) CFS agg, staggered laws	Kesiddai aggregation	0	(.014)	(.017)
В.	AR(1) DISTRIBUTION		(.014)	(.017)
17) $AR(1), \rho = .8$	Simple aggregation	50	.050	.243
	-1		(.013)	(.025)
18) $AR(1)$, $\rho = .8$	Residual aggregation	50	.045	.235
			(.012)	(.024)
19) AR(1), ρ = .8, staggered laws	Residual aggregation	50	.075	.355
			(.015)	(.028)

Solutions: "Cluster" within states (over time)

- simple, easy to implement
- Works well for N=10
- But this is only one data set and one variable (CPS, log weekly earnings)

TABLE VIII
ARBITRARY VARIANCE-COVARIANCE MATRIX

			Rejection rate				
Data	Technique	N	No effect	2% effect			
	A. CPS DATA						
1) CPS aggregate	OLS	50	.49	.663			
			(.025)	(.024)			
CPS aggregate	Cluster	50	.063	.268			
			(.012)	(.022)			
CPS aggregate	OLS	20	.385	.535			
			(.024)	(.025)			
4) CPS aggregate	Cluster	20	.058	.13			
			(.011)	(.017)			
CPS aggregate	old	10	.443	.51			
			(.025)	(.025)			
6) CPS aggregate	Cluster	10	.08	.12			
			(.014)	(.016)			
7) CPS aggregate	OLS	6	.383	.433			
			(.024)	(.025)			
8) CPS aggregate	Cluster	6	.115	.118			
			(.016)	(.016)			
B. AR(1) DISTRIBUTION							
9) AR(1), $\rho = .8$	Cluster	50	.045	.275			
			(.012)	(.026)			
10) AR(1), $\rho = 0$	Cluster	50	.035	.74			
			(.011)	(.025)			

Current Standard Practice

- Be conservative: cluster by group or time (not the interaction) and report the larger std. error
 - note: this may get size and power wrong
- Better.. you can cluster on both! Cameron, Gelbach, and Miller (2006, NBER Technical WP) method not coded in Stata yet, but you can get an .ado from Doug Miller's Stata page http://www.econ.ucdavis.edu/faculty/dlmiller/statafiles/
- Do you have enough groups for a normal approximation?
 Check with a "Wild Bootstrap" Cameron, Gelbach, Miller (ReStat 2008);
 - .do file on Miller's page.
- May be argument for using Newey-West std. errors.
- Ask Gordon Dahl, who is working on a better method

Exam?

Wed Dec 7 in Granger room, 3PM